Bridging convex and nonconvex optimization in noisy matrix completion: Stability and uncertainty quantification



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## Convex relaxation for low-rank structure

$$\begin{array}{ll} \mbox{minimize} & \|\boldsymbol{Z}\|_* \triangleq \sum_i \sigma_i(\boldsymbol{Z}) \\ \mbox{subj. to} & \mbox{noiseless data constraints} \end{array}$$



low-rank matrix Composition C by Piet Mondrian



semidefinite relaxation

### Convex relaxation for low-rank structure

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matrix sensing (Recht, Fazel, Parrilo '07)

- phase retrieval
  - matrix completion
  - robust PCA

. . .

- Hankel matrix completion  $\checkmark$
- blind deconvolution
- joint alignment / matching

(Candès, Strohmer, Voroninski '11, Candès, Li '12) (Candès, Recht '08, Candès, Tao '08, Gross '09) (Chandrasekaran et al. '09, Candès et al. '09) (Fazel et al. '13, Chen, Chi '13, Cai et al. '15) (Ahmed, Recht, Romberg '12, Ling, Strohmer '15)

(Chen, Huang, Guibas'14)

 $\begin{array}{ll} \underset{Z}{\text{minimize}} & \|Z\|_{*} \\ \text{subj. to} & \text{noisy data constraints} \end{array}$ 



semidefinite relaxation

low-rank matrix Composition C by Piet Mondrian



$$\underset{\boldsymbol{Z}}{\text{minimize}} \qquad \underbrace{f(\boldsymbol{Z}; \mathsf{data})}_{\text{empirical loss}} + \lambda \|\boldsymbol{Z}\|_{*}$$

- ✓ matrix sensing (RIP measurements) (Candès, Plan '10)
- ✓ phase retrieval (Gaussian measurements)
- ? matrix completion
  - (Candès, Plan '09, Negahban, Wainwright '10, Koltchinskii et al. '10)
- ? robust PCA
- ? Hankel matrix completion
- blind deconvolution (Ahmed, Recht, Romberg '12, Ling, Strohmer '15)
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(Candès et al. '11)

(Chen, Chi'13)

(Zhou, Li, Wright, Candès, Ma'10)

- ✓ matrix sensing (RIP measurements) (Candès, Plan '10)
- ✓ phase retrieval (Gaussian measurements)

(Candès et al. '11)

(Chen, Chi'13)

? this talk: matrix completion

(Candès, Plan '09, Negahban, Wainwright '10, Koltchinskii et al. '10)

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. . .

## Low-rank matrix completion



figure credit: E. J. Candès

Given partial samples of a low-rank matrix  $M^{\star}$ , fill in missing entries



recommendation systems







channel estimation

## Noisy low-rank matrix completion





$$\begin{bmatrix} \checkmark & ? & ? & ? & \checkmark & ? \\ ? & ? & \checkmark & \checkmark & ? & ? \\ \checkmark & ? & ? & \checkmark & \checkmark & ? & ? \\ \checkmark & ? & ? & \checkmark & ? & ? & ? \\ ? & ? & \checkmark & ? & ? & ? & \checkmark \\ \checkmark & ? & ? & ? & ? & ? & ? \\ ? & \checkmark & ? & ? & \checkmark & ? & ? \\ ? & ? & \checkmark & \checkmark & ? & ? \\ ? & ? & \checkmark & \checkmark & ? & ? \\ \end{bmatrix}$$

unknown rank-r matrix  $oldsymbol{M}^{\star} \in \mathbb{R}^{n imes n}$ 

sampling set  $\Omega$ 

## Noisy low-rank matrix completion

 $\begin{array}{ll} \text{observations:} & M_{i,j} = M_{i,j}^{\star} + \text{noise}, \quad (i,j) \in \Omega\\ \text{goal:} & \text{estimate } \boldsymbol{M}^{\star} \end{array}$ 



- random sampling: each  $(i, j) \in \Omega$  with prob. p
- random noise: i.i.d. sub-Gaussian noise with variance  $\sigma^2$
- true matrix  $M^{\star} \in \mathbb{R}^{n \times n}$ : rank r = O(1), incoherent, ...







minimax limit	$\sigma \sqrt{n/p}$
Candès, Plan '09	$\sigma n^{1.5}$
Negahban, Wainwright '10	$\max\{\sigma, \ \boldsymbol{M}^{\star}\ _{\infty}\}\sqrt{n/p}$
Koltchinskii, Tsybakov, Lounici '10	$\max\{\sigma, \ \boldsymbol{M}^{\star}\ _{\infty}\}\sqrt{n/p}$



#### Matrix Completion with Noise

Emmanuel J. Candès and Yaniv Plan



Existing theory for convex relaxation does not match practice ....

#### Matrix Completion with Noise

Emmanuel J. Candès and Yaniv Plan

with adversarial noise. Consequently, our analysis looses a  $\sqrt{n}$  factor vis a vis an optimal bound that is achievable via the help of an oracle.

Existing theory for convex relaxation does not match practice ....

Strategy:  $M^{\text{cvx}}$  is optimizer if there exists W s.t. dual certificate

 $(\boldsymbol{M}^{\mathsf{cvx}}, \boldsymbol{W})$  obeys KKT optimality condition

Strategy:  $M^{cvx}$  is optimizer if there exists W s.t.

 $(\boldsymbol{M}^{\mathsf{cvx}}, \boldsymbol{W})$  obeys KKT optimality condition



David Gross

• noiseless case:  $\underbrace{M^{\mathsf{cvx}} \leftarrow M^{\star}}_{\mathsf{exact recovery}}$ ;  $W \leftarrow \mathsf{golfing scheme}$ 

Strategy:  $M^{cvx}$  is optimizer if there exists W s.t.

 $(\boldsymbol{M}^{\mathsf{cvx}}, \boldsymbol{W})$  obeys KKT optimality condition



David Gross

- noiseless case:  $\underbrace{M^{\mathsf{cvx}} \leftarrow M^{\star}}_{\mathsf{exact recovery}}$ ;  $W \leftarrow \mathsf{golfing scheme}$
- noisy case:  $M^{\mathsf{cvx}}$  is very complicated, hard to construct W ...







#### nonconvex optimization

## A detour: nonconvex optimization

**Burer–Monteiro:** represent Z by  $XY^{\top}$  with  $X, Y \in \mathbb{R}^{n \times r}$ 

low-rank factors



## A detour: nonconvex optimization

**Burer–Monteiro:** represent Z by  $XY^{\top}$  with  $X, Y \in \mathbb{R}^{n \times r}$ 

low-rank factors



$$\underset{\boldsymbol{X},\boldsymbol{Y} \in \mathbb{R}^{n \times r}}{\text{minimize}} \quad f(\boldsymbol{X},\boldsymbol{Y}) = \underbrace{\sum_{(i,j) \in \Omega} \left[ \left( \boldsymbol{X} \boldsymbol{Y}^{\top} \right)_{i,j} - M_{i,j} \right]^2}_{\text{squared loss}} + \operatorname{reg}(\boldsymbol{X},\boldsymbol{Y})$$

# A detour: nonconvex optimization

- Burer, Monteiro '03
- Rennie, Srebro '05
- Keshavan, Montanari, Oh'09'10
- Jain, Netrapalli, Sanghavi '12
- Hardt '13
- Sun, Luo'14
- Chen, Wainwright '15
- Tu, Boczar, Simchowitz, Soltanolkotabi, Recht'15
- Zhao, Wang, Liu '15
- Zheng, Lafferty '16
- Yi, Park, Chen, Caramanis'16
- Ge, Lee, Ma'16
- Ge, Jin, Zheng '17
- Ma, Wang, Chi, Chen '17
- Chen, Li '18
- Chen, Liu, Li'19
- ...

$$\min_{\boldsymbol{X},\boldsymbol{Y} \in \mathbb{R}^{n \times r} } f(\boldsymbol{X},\boldsymbol{Y}) = \sum_{(i,j) \in \Omega} \left[ \left( \boldsymbol{X} \boldsymbol{Y}^\top \right)_{i,j} - M_{i,j} \right]^2 + \operatorname{reg}(\boldsymbol{X},\boldsymbol{Y})$$



- suitable initialization:  $({oldsymbol X}^0, {oldsymbol Y}^0)$
- gradient descent: for  $t = 0, 1, \ldots$

$$\boldsymbol{X}^{t+1} = \boldsymbol{X}^t - \eta_t \, \nabla_{\boldsymbol{X}} f(\boldsymbol{X}^t, \boldsymbol{Y}^t)$$
$$\boldsymbol{Y}^{t+1} = \boldsymbol{Y}^t - \eta_t \, \nabla_{\boldsymbol{Y}} f(\boldsymbol{X}^t, \boldsymbol{Y}^t)$$

- random sampling: each  $(i, j) \in \Omega$  with prob. p
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- true matrix  $M^{\star} \in \mathbb{R}^{n \times n}$ : r = O(1), incoherent, ...









**convex:** minimize 
$$\sum_{\boldsymbol{Z} \in \mathbb{R}^{n imes n}} \sum_{(i,j) \in \Omega} \left( Z_{i,j} - M_{i,j} \right)^2 + \lambda \|\boldsymbol{Z}\|_*$$

nonconvex: minimize  

$$\mathbf{X}, \mathbf{Y} \in \mathbb{R}^{n \times r} \sum_{(i,j) \in \Omega} \left[ \left( \mathbf{X} \mathbf{Y}^{\top} \right)_{i,j} - M_{i,j} \right]^2 + \underbrace{\frac{\lambda}{2} \| \mathbf{X} \|_{\mathrm{F}}^2 + \frac{\lambda}{2} \| \mathbf{Y} \|_{\mathrm{F}}^2}_{\mathsf{reg}(\mathbf{X}, \mathbf{Y})}$$

$$- \|Z\|_* = \min_{Z=XY^{\top}} \frac{1}{2} \|X\|_{\rm F}^2 + \frac{1}{2} \|Y\|_{\rm F}^2$$

### A motivating experiment

$$n = 1000, r = 5, p = 0.2, \lambda = 5\sigma\sqrt{np}$$



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$$n = 1000, r = 5, p = 0.2, \lambda = 5\sigma\sqrt{np}$$



Convex and nonconvex solutions are exceedingly close!


- random sampling: each  $(i,j) \in \Omega$  with prob.  $p \gtrsim \frac{\log^3 n}{n}$
- random noise: i.i.d. sub-Gaussian with variance  $\sigma^2$  (not too large)
- true matrix  $M^{\star} \in \mathbb{R}^{n \times n}$ : r = O(1), incoherent, well-conditioned

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$$\underset{\boldsymbol{Z} \in \mathbb{R}^{n \times n}}{\text{minimize}} \quad \sum_{(i,j) \in \Omega} \left( Z_{i,j} - M_{i,j} \right)^2 + \lambda \|\boldsymbol{Z}\|_* \qquad (\lambda \asymp \sigma \sqrt{np})$$

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#### Theorem 1 (Chen, Chi, Fan, Ma, Yan '19)

With high prob., any minimizer  $M^{cvx}$  of convex program obeys 1.  $M^{cvx}$  is nearly rank-r

$$\left\| \boldsymbol{M}^{\mathsf{cvx}} - \mathsf{proj}_{\mathit{rank-r}}(\boldsymbol{M}^{\mathsf{cvx}}) 
ight\|_{\mathrm{F}} \ll rac{1}{n^5} \cdot \sigma \sqrt{rac{n}{p}}$$

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2. 
$$\|\boldsymbol{M}^{\mathsf{cvx}} - \boldsymbol{M}^{\star}\|_{\mathrm{F}} \lesssim \sigma \sqrt{\frac{n}{p}}$$

- random sampling: each  $(i,j) \in \Omega$  with prob.  $p \gtrsim \frac{\log^3 n}{n}$
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$$\min_{\boldsymbol{Z} \in \mathbb{R}^{n \times n}} \sum_{(i,j) \in \Omega} \left( Z_{i,j} - M_{i,j} \right)^2 + \lambda \|\boldsymbol{Z}\|_* \qquad (\lambda \asymp \sigma \sqrt{np})$$

#### Theorem 1 (Chen, Chi, Fan, Ma, Yan '19)

With high prob., any minimizer  $M^{cvx}$  of convex program obeys 1.  $M^{cvx}$  is nearly rank-r 2.  $\|M^{cvx} - M^{\star}\|_{F} \lesssim \sigma \sqrt{\frac{n}{p}}$  $\|M^{cvx} - M^{\star}\|_{\infty} \lesssim \sigma \sqrt{\frac{n \log n}{p}} \cdot \frac{1}{n}$ 

$$ig\|oldsymbol{M}^{\mathsf{cvx}}-oldsymbol{M}^{\star}ig\|_{\mathrm{F}}\ \lesssim\sigma\sqrt{rac{n}{p}}$$



• minimax optimal when r = O(1)

$$\|M^{\text{cvx}} - M^{\star}\|_{\text{F}} \lesssim \sigma \sqrt{\frac{n}{p}} \qquad \|M^{\text{cvx}} - M^{\star}\|_{\infty} \lesssim \sigma \sqrt{\frac{n \log n}{p}} \cdot \frac{1}{n}$$

- minimax optimal when  $\boldsymbol{r}=\boldsymbol{O}(1)$
- estimation errors are spread out across all entries

#### No need to enforce spikiness constraint as in Negahban & Wainwright

$$\min_{\|\boldsymbol{Z}\|_{\infty} \leq \alpha} \quad \sum_{(i,j) \in \Omega} \left( Z_{i,j} - M_{i,j} \right)^2 + \lambda \|\boldsymbol{Z}\|_* \qquad (\text{Negahban et al.})$$

• convex programming automatically controls spikiness of solutions

# Statistical guarantees for iterative algorithms



Many algorithms (e.g. SVT, SOFT-IMPUTE, FPC, FISTA) have been proposed to solve (1), typically without statistical guarantees

# Statistical guarantees for iterative algorithms



Many algorithms (e.g. SVT, SOFT-IMPUTE, FPC, FISTA) have been proposed to solve (1), typically without statistical guarantees

We provide statistical guarantees for any  ${\pmb Z}$  with  $g({\pmb Z}) \leq g({\pmb Z}_{\sf opt}) + \varepsilon$  for some sufficiently small  $\varepsilon > 0$ 

- random sampling: each  $(i, j) \in \Omega$  with prob.  $p \gtrsim \frac{r^2 \log^3 n}{n}$
- random noise: i.i.d. sub-Gaussian with variance  $\sigma^2$  (not too large)
- true matrix  $M^{\star} \in \mathbb{R}^{n imes n}$ : incoherent, well-conditioned

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#### Theorem 2 (Chen, Chi, Fan, Ma, Yan'19)

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2. 
$$\|\boldsymbol{M}^{\mathsf{cvx}} - \boldsymbol{M}^{\star}\|_{\mathrm{F}} \lesssim \frac{\sigma}{\sigma_{\min}(\boldsymbol{M}^{\star})} \sqrt{\frac{n}{p}} \|\boldsymbol{M}^{\star}\|_{\mathrm{F}} \\ \|\boldsymbol{M}^{\mathsf{cvx}} - \boldsymbol{M}^{\star}\|_{\infty} \lesssim \sqrt{r} \frac{\sigma}{\sigma_{\min}(\boldsymbol{M}^{\star})} \sqrt{\frac{n\log n}{p}} \|\boldsymbol{M}^{\star}\|_{\infty} \\ \|\boldsymbol{M}^{\mathsf{cvx}} - \boldsymbol{M}^{\star}\| \lesssim \frac{\sigma}{\sigma_{\min}(\boldsymbol{M}^{\star})} \sqrt{\frac{n}{p}} \|\boldsymbol{M}^{\star}\|$$

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sample complexity bound  $O(nr^2 \log^3 n)$  is suboptimal in r!

*A little analysis:* connection between convex and nonconvex solutions

# Link between convex and nonconvex optimizers

 $(\pmb{X}, \pmb{Y})$  is nonconvex optimizer

# Link between convex and nonconvex optimizers

(X,Y) is nonconvex optimizer  $\stackrel{?}{\Longrightarrow} XY^{ op}$  is convex solution

# Link between convex and nonconvex optimizers

- (X,Y) is close to truth (in  $\ell_{2,\infty}$  sense)
- a little condition on noise size



### (X,Y) is nonconvex optimizer $\stackrel{\checkmark}{\Longrightarrow} XY^{ op}$ is convex solution



Issue: we do NOT know properties of nonconvex optimizers

• It is unclear whether nonconvex algorithms converge to optimizers (due to lack of strong convexity)

**Strategy:** resort to "approximate stationary points" instead  $\nabla f(\mathbf{X}, \mathbf{Y}) \approx \mathbf{0}$ 

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starting from  $(X^0, Y^0) =$ truth or spectral initialization:

$$\begin{aligned} \mathbf{X}^{t+1} &= \mathbf{X}^t - \eta \, \nabla_{\mathbf{X}} f(\mathbf{X}^t, \mathbf{Y}^t) \\ \mathbf{Y}^{t+1} &= \mathbf{Y}^t - \eta \, \nabla_{\mathbf{Y}} f(\mathbf{X}^t, \mathbf{Y}^t) \end{aligned} \qquad t = 0, 1, \cdots, T \end{aligned}$$

**Strategy:** resort to "approximate stationary points" instead  $\nabla f(\mathbf{X}, \mathbf{Y}) \approx \mathbf{0}$ 

starting from  $(\boldsymbol{X}^0, \boldsymbol{Y}^0) = \mathsf{truth}$  or spectral initialization:

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- $\bullet\,$  when T is large: there exists point with very small gradient
- hopefully not far from  $({\pmb{X}}^\star, {\pmb{Y}}^\star)$

 $\|\nabla f(\boldsymbol{X},\boldsymbol{Y})\|_{\mathrm{F}} \lesssim \frac{1}{\sqrt{nT}}$ 

# Analyzing nonconvex GD: leave-one-out analysis

Leave out a small amount of information from data and run GD

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Leave out a small amount of information from data and run GD

- Stein '72
- El Karoui, Bean, Bickel, Lim, Yu'13
- El Karoui '15
- Javanmard, Montanari '15
- Zhong, Boumal '17
- Lei, Bickel, El Karoui '17
- Sur, Chen, Candès'17
- Abbe, Fan, Wang, Zhong '17
- Chen, Fan, Ma, Wang'17
- Ma, Wang, Chi, Chen '17
- Chen, Chi, Fan, Ma'18
- Ding, Chen '18
- Dong, Shi'18
- Chen, Liu, Li'19

# Analyzing nonconvex GD: leave-one-out analysis

For each  $1 \le l \le n$ , introduce leave-one-out iterates  $X^{t,(l)}$  by replacing  $l^{th}$  row and column with true values



- exploit partial statistical independence
- exploit leave-one-out stability

Inference and uncertainty quantification

# Reasoning about uncertainty



### Reasoning about uncertainty



matrix completion

3	2	4	2	1
4	2	6	4	2
3	1	5	4	2
3	1	4	3	1
1	0	3	3	2

### Reasoning about uncertainty



How to assess uncertainty, or "confidence", of obtained estimates?

#### INFERENCE IN HIGH DIMENSIONAL REGRESSION

organized by Peter Buehlmann, Andrea Montanari, and Jonathan Taylor



(3) <u>Confidence intervals for matrix completion</u>. In matrix completion, the data analyst is given a large data matrix with a number of missing entries. In many interesting applications (e.g. to collaborative filtering) it is indeed the case that the vast majority of entries is missing. In order to fill the missing entries, the assumption is made that the underlying –unknown– matrix has a low-rank structure.

Substantial work has been devoted to methods for computing point estimates of the missing entries. In applications, it would be very interesting to compute confidence intervals as well. This requires developing distributional characterizations of standard matrix completion methods.

$$\boldsymbol{M}^{\mathsf{cvx}} \triangleq \operatorname*{arg\,min}_{\boldsymbol{Z} \in \mathbb{R}^{n \times n}} \sum_{(i,j) \in \Omega} \left( Z_{i,j} - M_{i,j} \right)^2 + \lambda \|\boldsymbol{Z}\|_*$$

- convex estimate  $M^{ ext{cvx}}$  is biased towards small norm

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- very challenging to pin down distributions of obtained estimates

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- convex estimate  $M^{ ext{cvx}}$  is biased towards small norm
- very challenging to pin down distributions of obtained estimates
- existing orderwise bounds come with unspecified (but huge) pre-constants

 $\longrightarrow$  overly wide confidence intervals

- inspired by Zhang, Zhang '11, van de Geer et al. '13, Javanmard, Montanari '13



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$$M^{\mathsf{cvx}} \stackrel{\mathsf{de-biasing}}{\longrightarrow} \underbrace{M^{\mathsf{cvx}} + rac{1}{p} \mathcal{P}_{\Omega}(M^{\star} + E - M^{\mathsf{cvx}})}_{(\mathsf{nearly}) \, \mathsf{unbiased \, estimate \, of} \, M^{\star}}$$

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• issue: high-rank after de-biasing; statistical accuracy suffers

$$M^{\mathsf{cvx}} \xrightarrow{\mathsf{de-biasing}} \underbrace{\operatorname{proj}_{\mathsf{rank-}r} \left( M^{\mathsf{cvx}} + \frac{1}{p} \mathcal{P}_{\Omega}(M^{\star} + E - M^{\mathsf{cvx}}) \right)}_{1 \text{ iteration of singular value projection (Jain, Meka, Dhillon '10)}} =: M^{\mathrm{d}}$$

- issue: high-rank after de-biasing; statistical accuracy suffers
- solution: low-rank projection



## Distributional guarantees for low-rank factors

- random sampling: each  $(i,j) \in \Omega$  with prob.  $p \gtrsim \frac{\log^3 n}{n}$
- random noise: i.i.d.  $\mathcal{N}(0,\sigma^2)$  (not too large)
- true matrix  $M^{\star} \in \mathbb{R}^{n \times n}$ : r = O(1), incoherent, well-conditioned
- regularization parameter:  $\lambda \asymp \sigma \sqrt{np}$

$$X^{\mathsf{d}}Y^{\mathsf{d}^{\top}} \leftarrow \underbrace{\text{balanced}}_{X^{\mathsf{d}^{\top}}X^{\mathsf{d}}=Y^{\mathsf{d}^{\top}}Y^{\mathsf{d}}}$$
 rank- $r$  decomp. of  $M^{\mathsf{d}}$   
 $X^{\star}Y^{\star^{\top}} \leftarrow \underbrace{\text{balanced}}_{X^{\star^{\top}}X^{\star}=Y^{\star^{\top}}Y^{\star}}$  rank- $r$  decomp. of  $M^{\star}$ 

## Distributional guarantees for low-rank factors

$$\begin{array}{l} \boldsymbol{X}^{\mathsf{d}}\boldsymbol{Y}^{\mathsf{d}^{\top}} \leftarrow \underbrace{\text{balanced}}_{\boldsymbol{X}^{\mathsf{d}^{\top}}\boldsymbol{X}^{\mathsf{d}}=\boldsymbol{Y}^{\mathsf{d}^{\top}}\boldsymbol{Y}^{\mathsf{d}}} \text{ rank-}r \text{ approx. of } \boldsymbol{M}^{\mathsf{d}} \\ \boldsymbol{X}^{\star}\boldsymbol{Y}^{\star^{\top}} \leftarrow \underbrace{\text{balanced}}_{\boldsymbol{X}^{\star^{\top}}\boldsymbol{X}^{\star}=\boldsymbol{Y}^{\star^{\top}}\boldsymbol{Y}^{\star}} \text{ rank-}r \text{ decomp. of } \boldsymbol{M}^{\star} \end{array}$$

#### Theorem 3 (Chen, Fan, Ma, Yan '19)

With high prob., there exists global rotation matrix  $\boldsymbol{H} \in \mathbb{R}^{r \times r}$  s.t.  $\boldsymbol{X}^{\mathrm{d}}\boldsymbol{H} - \boldsymbol{X}^{\star} \approx \boldsymbol{Z}^{X}, \qquad \boldsymbol{Z}_{i,\cdot}^{X} \stackrel{\mathrm{ind.}}{\sim} \mathcal{N}(\boldsymbol{0}, \frac{\sigma^{2}}{p} (\boldsymbol{Y}^{\star \top} \boldsymbol{Y}^{\star})^{-1})$  $\boldsymbol{Y}^{\mathrm{d}}\boldsymbol{H} - \boldsymbol{Y}^{\star} \approx \boldsymbol{Z}^{Y}, \qquad \boldsymbol{Z}_{i,\cdot}^{Y} \stackrel{\mathrm{ind.}}{\sim} \mathcal{N}(\boldsymbol{0}, \frac{\sigma^{2}}{p} (\boldsymbol{X}^{\star \top} \boldsymbol{X}^{\star})^{-1})$ 

$$\begin{split} \boldsymbol{X}^{\mathrm{d}}\boldsymbol{H} - \boldsymbol{X}^{\star} &\approx \boldsymbol{Z}^{X}, \qquad \boldsymbol{Z}_{i,\cdot}^{X} \stackrel{\mathrm{ind.}}{\sim} \mathcal{N}(\boldsymbol{0}, \frac{\sigma^{2}}{p} (\boldsymbol{Y}^{\star \top} \boldsymbol{Y}^{\star})^{-1}) \\ \boldsymbol{Y}^{\mathrm{d}}\boldsymbol{H} - \boldsymbol{Y}^{\star} &\approx \boldsymbol{Z}^{Y}, \qquad \boldsymbol{Z}_{i,\cdot}^{Y} \stackrel{\mathrm{ind.}}{\sim} \mathcal{N}(\boldsymbol{0}, \frac{\sigma^{2}}{p} (\boldsymbol{X}^{\star \top} \boldsymbol{X}^{\star})^{-1}) \end{split}$$

• estimation errors for different rows of  $X^{\star}$  are nearly independent

$$oldsymbol{X}^{\mathrm{d}}_{i,\cdot}oldsymbol{H} - oldsymbol{X}^{\star}_{i,\cdot}$$
 nearly ind. of  $oldsymbol{X}^{\mathrm{d}}_{j,\cdot}oldsymbol{H} - oldsymbol{X}^{\star}_{j,\cdot}$ 

$$\begin{split} & \boldsymbol{X}^{\mathrm{d}}\boldsymbol{H} - \boldsymbol{X}^{\star} \; \approx \; \boldsymbol{Z}^{X}, \qquad \boldsymbol{Z}^{X}_{i,\cdot} \stackrel{\mathrm{ind.}}{\sim} \mathcal{N}(\boldsymbol{0}, \frac{\sigma^{2}}{p} (\boldsymbol{Y}^{\star\top} \boldsymbol{Y}^{\star})^{-1}) \\ & \boldsymbol{Y}^{\mathrm{d}}\boldsymbol{H} - \boldsymbol{Y}^{\star} \; \approx \; \boldsymbol{Z}^{Y}, \qquad \boldsymbol{Z}^{Y}_{i,\cdot} \stackrel{\mathrm{ind.}}{\sim} \mathcal{N}(\boldsymbol{0}, \frac{\sigma^{2}}{p} (\boldsymbol{X}^{\star\top} \boldsymbol{X}^{\star})^{-1}) \end{split}$$

• accurate uncertainty quantification for low-rank factors, e.g.

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• accurate uncertainty quantification for matrix entries: if  $\|X_{i,\cdot}^{\star}\|_2 + \|Y_{j,\cdot}^{\star}\|_2$  is not too small, then

 $M_{i,j}^{\mathsf{d}} - M_{i,j}^{\star} \sim \mathcal{N}(0, v_{i,j}^{\star}) + \text{negligible term}$ 

where 
$$v_{i,j}^{\star} \triangleq \frac{\sigma^2}{p} \Big\{ \boldsymbol{X}_{i,\cdot}^{\star} (\boldsymbol{X}^{\star \top} \boldsymbol{X}^{\star})^{-1} \boldsymbol{X}_{i,\cdot}^{\star \top} + \boldsymbol{Y}_{j,\cdot}^{\star} (\boldsymbol{Y}^{\star \top} \boldsymbol{Y}^{\star})^{-1} \boldsymbol{Y}_{j,\cdot}^{\star \top} \Big\}$$

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$$n = 1000, p = 0.2, r = 5, ||\mathbf{M}^{\star}|| = 1, \kappa = 1, \sigma = 10^{-3}$$









Same inference procedures work for both cvx & noncvx estimates!

Consider rank-1 PSD case  $M^{\star} = x^{\star}x^{\star \top}$ , p = 1 (no missing data)

$$\mathsf{minimize}_{oldsymbol{x}} \qquad rac{1}{2} \|oldsymbol{x}oldsymbol{x}^ op - oldsymbol{x}^\staroldsymbol{\pi}^ op - oldsymbol{E}\|_{\mathrm{F}}^2 + \lambda \|oldsymbol{x}\|_2^2$$

Consider rank-1 PSD case  $M^{\star} = x^{\star}x^{\star op}$ , p = 1 (no missing data)

$$\mathsf{minimize}_{oldsymbol{x}} \qquad rac{1}{2} \|oldsymbol{x}oldsymbol{x}^ op - oldsymbol{x}^\staroldsymbol{x}^ op - oldsymbol{E}\|_{\mathrm{F}}^2 + \lambda \|oldsymbol{x}\|_2^2$$

• first-order optimality condition

$$(\boldsymbol{x}\boldsymbol{x}^{ op}-\boldsymbol{x}^{\star}\boldsymbol{x}^{\star op}-\boldsymbol{E})\boldsymbol{x}+\lambda \boldsymbol{x}=\boldsymbol{0}$$

$$(xx^{ op} - x^{\star}x^{\star op} - E)x \underbrace{+\lambda x}_{\mathsf{causes bias}} = \mathbf{0}$$

$$(oldsymbol{x}oldsymbol{x}^ op - oldsymbol{x}^\staroldsymbol{x}^{ op} - oldsymbol{E})oldsymbol{x} \stackrel{}{ op} = oldsymbol{0}$$
 causes bias

$$(oldsymbol{x}^{\mathrm{d} op}-oldsymbol{x}^{\star}oldsymbol{x}^{\star op}-oldsymbol{E})oldsymbol{x}^{\mathrm{d}}=oldsymbol{0},\qquadoldsymbol{x}^{\mathrm{d}}=\sqrt{rac{\lambda+\|oldsymbol{x}\|_2^2}{\|oldsymbol{x}\|_2^2}}\,oldsymbol{x}$$

$$(\boldsymbol{x}\boldsymbol{x}^{ op} - \boldsymbol{x}^{\star}\boldsymbol{x}^{\star op} - \boldsymbol{E})\boldsymbol{x} \underbrace{+\lambda \boldsymbol{x}}_{\mathsf{causes bias}} = \mathbf{0}$$

$$(\boldsymbol{x}^{\mathrm{d}}\boldsymbol{x}^{\mathrm{d} op} - \boldsymbol{x}^{\star}\boldsymbol{x}^{\star op} - \boldsymbol{E})\boldsymbol{x}^{\mathrm{d}} = \mathbf{0}, \qquad \boldsymbol{x}^{\mathrm{d}} = \sqrt{\frac{\lambda + \|\boldsymbol{x}\|_{2}^{2}}{\|\boldsymbol{x}\|_{2}^{2}}} \boldsymbol{x}$$

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$$(\boldsymbol{x}^{\mathrm{d}} - \boldsymbol{x}^{\star} = \underbrace{\frac{1}{\|\boldsymbol{x}^{\mathrm{d}}\|_{2}^{2}} \boldsymbol{E} \boldsymbol{x}^{\mathrm{d}}}_{\mathsf{nearly Gaussian}} + \underbrace{\frac{(\boldsymbol{x}^{\star} - \boldsymbol{x}^{\mathrm{d}})^{\top} \boldsymbol{x}^{\mathrm{d}}}{\|\boldsymbol{x}^{\mathrm{d}}\|_{2}^{2}} \boldsymbol{x}^{\star}}_{\mathsf{hopefully small}}$$

- inspired by Zhang, Zhang '11, van de Geer et al. '13, Javanmard, Montanari '13

$$oldsymbol{M}^{\mathsf{cvx}} riangleq rgmin_{oldsymbol{Z} \in \mathbb{R}^{n imes n}} \| \mathcal{P}_{\Omega}(oldsymbol{Z} - oldsymbol{M}) \|_{\mathrm{F}}^2 + \lambda \|oldsymbol{Z}\|_*$$

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$$oldsymbol{M}^{\mathsf{cvx}} riangleq rgmin_{oldsymbol{Z} \in \mathbb{R}^{n imes n}} egin{array}{c} \|\mathcal{P}_{\Omega}(oldsymbol{Z} - oldsymbol{M})\|_{\mathrm{F}}^2 + \lambda \|oldsymbol{Z}\|_* \ \end{array}$$

$$\widehat{\boldsymbol{M}}^{\mathrm{d}} = \boldsymbol{M}^{\mathsf{cvx}} + \mathsf{linear-map}\Big(\,\mathcal{P}_{\Omega}(\boldsymbol{M} - \boldsymbol{M}^{\mathsf{cvx}})\,\Big)$$

∜

- inspired by Zhang, Zhang '11, van de Geer et al. '13, Javanmard, Montanari '13

$$egin{aligned} oldsymbol{M}^{\mathsf{cvx}} & & ext{arg\,min} \quad \left\| \mathcal{P}_\Omega(oldsymbol{Z} - oldsymbol{M}) 
ight\|_{\mathrm{F}}^2 + \lambda \|oldsymbol{Z}\|_* \ & & \downarrow \ & \ & \widehat{oldsymbol{M}}^{\mathrm{d}} = oldsymbol{M}^{\mathsf{cvx}} + rac{1}{p} \mathcal{P}_T \Big( \left. \mathcal{P}_\Omegaig(oldsymbol{M} - oldsymbol{M}^{\mathsf{cvx}}) \, \Big) \end{aligned}$$

—  $\mathcal{P}_T:$  projection onto T (tangent space at  $\text{proj}_{\text{rank-r}}(\boldsymbol{M}^{\text{cvx}}))$ 

$$\widehat{M}^{\mathrm{d}}~pprox~M^{\mathrm{d}}$$
 !

#### Back to estimation: de-biased estimator is optimal

Distributional theory in turn allows us to track estimation accuracy

Distributional theory in turn allows us to track estimation accuracy



Distributional theory in turn allows us to track estimation accuracy



- precise characterization of estimation accuracy
- achieves full statistical efficiency (including pre-constant)

## Numerical evidence (r = 5, p = 0.2, n = 1000)



Euclidean estimation error vs. noise standard deviation  $\sigma$ 

## **Concluding remarks**



# **Concluding remarks**



• more general sampling patterns



"Noisy matrix completion: understanding statistical guarantees for convex relaxation via nonconvex optimization," Y. Chen, Y. Chi, J. Fan, C. Ma, Y. Yan, 2019

"Inference and uncertainty quantification for noisy matrix completion," Y. Chen, J. Fan, C. Ma, Y. Yan, 2019

Backup slides: gradient descent for nonconvex matrix completion

#### Gradient descent for nonconvex matrix completion



$$\begin{array}{ll} \boldsymbol{X}^{t+1} &= \boldsymbol{X}^t - \eta \, \nabla_{\boldsymbol{X}} f(\boldsymbol{X}^t, \boldsymbol{Y}^t) \\ \boldsymbol{Y}^{t+1} &= \boldsymbol{Y}^t - \eta \, \nabla_{\boldsymbol{Y}} f(\boldsymbol{X}^t, \boldsymbol{Y}^t) \end{array}$$

Prior works analyze regularized GD

- not guaranteed to return small-gradient solutions
- no  $\ell_{2,\infty}$  error control

- Keshavan et al. '09, Sun, Luo '15, Chen, Wainwright '15, Zheng, Lafferty '16

#### Gradient descent for nonconvex matrix completion



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Our work and Chen et al. analyze vanilla GD

- regularization-free
- optimal  $\ell_{2,\infty}$  error control

— Ma, Wang, Chi, Chen '17, Chen, Liu, Li '19



Two standard conditions that enable geometric convergence of GD



Two standard conditions that enable geometric convergence of GD

• (local) restricted strong convexity


Two standard conditions that enable geometric convergence of GD

- (local) restricted strong convexity
- (local) smoothness

f is said to be  $\alpha\text{-strongly convex and }\beta\text{-smooth if}$ 

$$\mathbf{0} \preceq \alpha \mathbf{I} \preceq \nabla^2 f(\mathbf{X}) \preceq \beta \mathbf{I}, \qquad \forall \mathbf{X}$$

$$\ell_2$$
 error contraction: GD with  $\eta = 1/\beta$  obeys $\| \boldsymbol{X}^{t+1} - \boldsymbol{X}^{\star} \|_{\mathrm{F}} \leq \left( 1 - rac{lpha}{eta} 
ight) \| \boldsymbol{X}^t - \boldsymbol{X}^{\star} \|_{\mathrm{F}}$ 



• X is not far away from X<sup>\*</sup>



- X is not far away from X<sup>\*</sup>
- X is incoherent w.r.t. standard basis vectors (incoherence region)



 $\|\boldsymbol{e}_2^\top(\boldsymbol{X}-\boldsymbol{X}^\star)\|_2 \leq \epsilon \|\boldsymbol{X}^\star\|_{2,\infty} \quad \|\boldsymbol{e}_1^\top(\boldsymbol{X}-\boldsymbol{X}^\star)\|_2 \leq \epsilon \|\boldsymbol{X}^\star\|_{2,\infty}$ 

- X is not far away from  $X^{\star}$
- X is incoherent w.r.t. standard basis vectors (incoherence region)



• Generic optimization theory does NOT ensure GD stays in incoherence region



• Generic optimization theory does NOT ensure GD stays in incoherence region



• Generic optimization theory does NOT ensure GD stays in incoherence region



- Generic optimization theory does NOT ensure GD stays in incoherence region
- Calls for new analysis tools

Leave out a small amount of information from data and run GD

Leave out a small amount of information from data and run GD

- Stein '72
- El Karoui, Bean, Bickel, Lim, Yu'13
- El Karoui '15
- Javanmard, Montanari '15
- Zhong, Boumal'17
- Lei, Bickel, El Karoui '17
- Sur, Chen, Candès'17
- Abbe, Fan, Wang, Zhong '17
- Chen, Fan, Ma, Wang'17
- Ma, Wang, Chi, Chen '17
- Chen, Chi, Fan, Ma'18
- Ding, Chen '18
- Dong, Shi'18
- Chen, Liu, Li'19

For each  $1 \le l \le n$ , introduce leave-one-out iterates  $X^{t,(l)}$  by replacing  $l^{th}$  row and column with true values





• Leave-one-out iterates  $\{X^{t,(l)}\}$  contains more information of  $l^{th}$  row of truth; indep. of randomness in  $l^{th}$  row



- Leave-one-out iterates  $\{X^{t,(l)}\}$  contains more information of  $l^{th}$  row of truth; indep. of randomness in  $l^{th}$  row
- Leave-one-out iterates  $\{ oldsymbol{X}^{t,(l)} \} pprox$  true iterates  $\{ oldsymbol{X}^t \}$