

## **Randomized linear algebra**



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# Outline

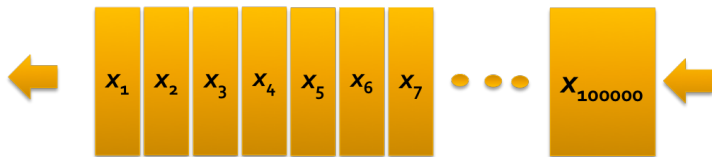
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- Approximate matrix multiplication
- Least squares approximation
- Low-rank matrix approximation
- Graph sparsification

**Main reference:** "*Lecture notes on randomized linear algebra,*"  
Michael W. Mahoney, 2016

# Efficient large-scale data processing

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When processing large-scale data (in particular, streaming data), we desire methods that can be performed with

- a few (e.g. one or two) passes of data
- limited memory (so impossible to store all data)
- low computational complexity

# Key idea: dimension reduction via random sketching

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- **random sampling:** randomly downsample data
  - often relies on information of data
- **random projection:** rotates / projects data onto lower dimensions
  - often data-agnostic

# **Approximate matrix multiplication**

# Matrix multiplication: a fundamental algebra task

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Given  $A \in \mathbb{R}^{m \times n}$  and  $B \in \mathbb{R}^{n \times p}$ , compute or approximate  $AB$

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**Algorithm 1.1** Vanilla algorithm for matrix multiplication

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```
1: for  $i = 1, \dots, m$  do  
2:   for  $k = 1, \dots, n$  do  
3:      $M_{i,k} = A_{i,:} B_{:,k}$   
4: return  $M$ 
```

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Computational complexity:  $O(mnp)$ , or  $O(n^3)$  if  $m = n = p$

For simplicity, we will assume  $m = n = p$  unless otherwise noted.

# Faster matrix multiplication?

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- **Strassen algorithms:** exact matrix multiplication
  - Computational complexity  $\approx O(n^{2.8})$
  - For various reasons, rarely used in practice
- Approximate solution?



# A simple randomized algorithm

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View  $AB$  as sum of rank-one matrices (or outer products)

$$AB = \sum_{i=1}^n A_{:,i} B_{i,:}$$

**Idea:** randomly sample  $r$  rank-one components

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**Algorithm 1.2** Basic randomized algorithm for matrix multiplication

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- 1: **for**  $l = 1, \dots, r$  **do**
- 2:   Pick  $i_l \in \{1, \dots, n\}$  i.i.d. with prob.  $\mathbb{P}\{i_l = k\} = p_k$
- 3: **return**

$$M = \sum_{l=1}^r \frac{1}{rp_{i_l}} A_{:,i_l} B_{i_l,:}$$

- 
- $\{p_k\}$ : importance sampling probabilities

# A simple randomized algorithm

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Rationale:  $M$  is *unbiased* estimate of  $AB$ , i.e.

$$\begin{aligned}\mathbb{E}[M] &= \sum_{l=1}^r \sum_k \mathbb{P}\{i_l = k\} \frac{1}{rp_k} \mathbf{A}_{:,k} \mathbf{B}_{k,:} \\ &= \sum_k \mathbf{A}_{:,k} \mathbf{B}_{k,:} = AB\end{aligned}$$

Clearly, approximation error (e.g.  $\|AB - M\|$ ) depends on  $\{p_k\}$ .

# Importance sampling probabilities

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- **Uniform sampling** ( $p_k \equiv \frac{1}{n}$ ): one can choose sampling set before looking at data, so it's implementable via one pass over data

Intuitively, one may prefer biasing towards larger rank-1 components

- **Nonuniform sampling**

$$p_k = \frac{\|\mathbf{A}_{:,k}\|_2 \|\mathbf{B}_{k,:}\|_2}{\sum_l \|\mathbf{A}_{:,l}\|_2 \|\mathbf{B}_{l,:}\|_2}$$

- $\{p_k\}$  can be computed using one pass and  $O(n)$  memory

# Optimal sampling probabilities?

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Let's measure approximation error by  $\mathbb{E} [\|M - AB\|_F^2]$ .

As it turns out,  $\mathbb{E} [\|M - AB\|_F^2]$  is minimized by

$$p_k = \frac{\|A_{:,k}\|_2 \|B_{k,:}\|_2}{\sum_l \|A_{:,l}\|_2 \|B_{l,:}\|_2} \quad (1.1)$$

Thus, we call (1.1) **optimal sampling probabilities** .

# Justification of optimal sampling probabilities

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Since  $\mathbb{E}[M] = AB$ , one has

$$\begin{aligned}\mathbb{E} [\|M - AB\|_F^2] &= \mathbb{E} \left[ \sum_{i,j} (M_{i,j} - \mathbf{A}_{i,:} \mathbf{B}_{:,j})^2 \right] = \sum_{i,j} \text{Var}[M_{i,j}] \\ &= \frac{1}{r} \sum_k \sum_{i,j} \frac{A_{i,k}^2 B_{k,j}^2}{p_k} - \frac{1}{r} \sum_{i,j} (\mathbf{A}_{i,:} \mathbf{B}_{:,j})^2 \quad (\text{check}) \\ &= \frac{1}{r} \sum_k \frac{1}{p_k} \|\mathbf{A}_{:,k}\|_2^2 \|\mathbf{B}_{k,:}\|_2^2 - \frac{1}{r} \|AB\|_F^2.\end{aligned}\quad (1.2)$$

In addition, Cauchy-Schwarz yields  $(\sum_k p_k) \left(\sum_k \frac{\alpha_k}{p_k}\right) \geq (\sum_k \sqrt{\alpha_k})^2$ , with equality attained if  $p_k \propto \sqrt{\alpha_k}$ . This implies

$$\mathbb{E} [\|M - AB\|_F^2] \geq \frac{1}{r} \left( \sum_k \|\mathbf{A}_{:,k}\|_2 \|\mathbf{B}_{k,:}\|_2 \right)^2 - \frac{1}{r} \|AB\|_F^2,$$

where lower bound is achieved when  $p_k \propto \|\mathbf{A}_{:,k}\|_2 \|\mathbf{B}_{k,:}\|_2$ .

# Error concentration

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Practically, one often hopes that approximation error is absolutely controlled most of the time. In other words, we desire a method whose estimate is sufficiently close to truth **with very high probability**

For approximate matrix multiplication, two error metrics are of particular interest

- Frobenius norm bound:  $\|M - AB\|_F$
- spectral norm bound:  $\|M - AB\|$

invoke **concentration of measure** results to control these errors

# Asymptotic notation

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- $f(n) \lesssim g(n)$  or  $f(n) = O(g(n))$  means

$$\lim_{n \rightarrow \infty} \frac{|f(n)|}{|g(n)|} \leq \text{const}$$

- $f(n) \gtrsim g(n)$  or  $f(n) = \Omega(g(n))$  means

$$\lim_{n \rightarrow \infty} \frac{|f(n)|}{|g(n)|} \geq \text{const}$$

- $f(n) \asymp g(n)$  or  $f(n) = \Theta(g(n))$  means

$$\text{const}_1 \leq \lim_{n \rightarrow \infty} \frac{|f(n)|}{|g(n)|} \leq \text{const}_2$$

- $f(n) = o(g(n))$  means

$$\lim_{n \rightarrow \infty} \frac{|f(n)|}{|g(n)|} = 0$$

# A hammer: matrix Bernstein inequality

## Theorem 1.1 (Matrix Bernstein inequality)

Let  $\{\mathbf{X}_l \in \mathbb{R}^{d_1 \times d_2}\}$  be a sequence of independent zero-mean random matrices. Assume each random matrix satisfies  $\|\mathbf{X}_l\| \leq R$ . Define  $V := \max \left\{ \left\| \mathbb{E} \left[ \sum_{l=1}^L \mathbf{X}_l \mathbf{X}_l^\top \right] \right\|, \left\| \mathbb{E} \left[ \sum_{l=1}^L \mathbf{X}_l^\top \mathbf{X}_l \right] \right\| \right\}$ . Then,

$$\mathbb{P} \left\{ \left\| \sum_{l=1}^L \mathbf{X}_l \right\| \geq \tau \right\} \leq (d_1 + d_2) \exp \left( \frac{-\tau^2/2}{V + R\tau/3} \right)$$

- *moderate-deviation regime* ( $\tau$  is not too large): sub-Gaussian tail behavior  $\exp(-\tau^2/V)$
- *large-deviation regime* ( $\tau$  is large): sub-exponential tail behavior  $\exp(-\tau/R)$  (slower decay)



# A hammer: matrix Bernstein inequality

## Theorem 1.1 (Matrix Bernstein inequality)

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$$\mathbb{P} \left\{ \left\| \sum_{l=1}^L \mathbf{X}_l \right\| \geq \tau \right\} \leq (d_1 + d_2) \exp \left( \frac{-\tau^2/2}{V + R\tau/3} \right)$$

- an alternative form (exercise): with prob.  $1 - O((d_1 + d_2)^{-10})$ ,

$$\left\| \sum_{l=1}^L \mathbf{X}_l \right\| \lesssim \sqrt{V \log(d_1 + d_2)} + R \log(d_1 + d_2)$$

# Frobenius norm error of matrix multiplication

## Theorem 1.2

Suppose  $p_k \geq \frac{\beta \|A_{:,k}\|_2 \|B_{k,:}\|_2}{\sum_l \|A_{:,l}\|_2 \|B_{l,:}\|_2}$  for some quantity  $0 < \beta \leq 1$ . If  $r \gtrsim \frac{\log n}{\beta}$ , then

$$\|M - AB\|_F \lesssim \sqrt{\frac{\log n}{\beta r}} \|A\|_F \|B\|_F$$

with prob. exceeding  $1 - O(n^{-10})$

## Proof of Theorem 1.2

Clearly,  $\text{vec}(M) = \sum_{l=1}^r \mathbf{X}_l$ , where

$\mathbf{X}_l = \sum_{k=1}^n \frac{1}{rp_k} \mathbf{A}_{:,k} \otimes \mathbf{B}_{k,:}^\top \mathbb{1}\{i_l = k\}$ . These matrices  $\{\mathbf{X}_l\}$  obey

$$\|\mathbf{X}_l\|_2 \leq \max_k \frac{1}{rp_k} \|\mathbf{A}_{:,k}\|_2 \|\mathbf{B}_{k,:}\|_2 \asymp \frac{1}{\beta r} \sum_{k=1}^n \|\mathbf{A}_{:,k}\|_2 \|\mathbf{B}_{k,:}\|_2 := R$$

$$\mathbb{E} \left[ \sum_{l=1}^r \|\mathbf{X}_l\|_2^2 \right] = r \sum_{k=1}^n \mathbb{P}\{i_l = k\} \frac{\|\mathbf{A}_{:,k}\|_2^2 \|\mathbf{B}_{k,:}\|_2^2}{r^2 p_k^2} \leq \underbrace{\frac{(\sum_{k=1}^n \|\mathbf{A}_{k,:}\|_2 \|\mathbf{B}_{k,:}\|_2)^2}{\beta r}}_{:=V}$$

Invoke matrix Bernstein to arrive at

$$\begin{aligned} \|\mathbf{M} - \mathbf{AB}\|_F &= \left\| \sum_{l=1}^r (\mathbf{X}_l - \mathbb{E}[\mathbf{X}_l]) \right\|_2 \lesssim \sqrt{V \log n} + R \log n \\ &\asymp \sqrt{\frac{\log n}{\beta r}} \left( \sum_{k=1}^n \|\mathbf{A}_{k,:}\|_2 \|\mathbf{B}_{k,:}\|_2 \right) \leq \sqrt{\frac{\log n}{\beta r}} \|\mathbf{A}\|_F \|\mathbf{B}\|_F \quad (\text{Cauchy-Schwarz}) \end{aligned}$$

# Spectral norm error of matrix multiplication

## Theorem 1.3

Suppose  $p_k \geq \frac{\beta \|\mathbf{A}_{:,k}\|_2^2}{\|\mathbf{A}\|_F^2}$  for some quantity  $0 < \beta \leq 1$ , and

$r \gtrsim \frac{\|\mathbf{A}\|_F^2}{\beta \|\mathbf{A}\|^2 \log n}$ . Then the estimate  $\mathbf{M}$  returned by Algorithm 1.2 obeys

$$\|\mathbf{M} - \mathbf{A}\mathbf{A}^\top\| \lesssim \sqrt{\frac{\log n}{\beta r}} \|\mathbf{A}\|_F \|\mathbf{A}\|$$

with prob. exceeding  $1 - O(n^{-10})$

- If  $r \gtrsim \underbrace{\frac{\|\mathbf{A}\|_F^2}{\|\mathbf{A}\|^2}}_{\text{stable rank}} \frac{\log n}{\varepsilon^2 \beta}$ , then  $\|\mathbf{M} - \mathbf{A}\mathbf{A}^\top\| \lesssim \varepsilon \|\mathbf{A}\|^2$

- can be generalized to approximate  $\mathbf{A}\mathbf{B}$  (Magen, Zouzias '11)

## Proof of Theorem 1.3

Write  $M = \sum_{l=1}^r Z_l$ , where  $Z_l = \sum_{k=1}^n \frac{1}{rp_k} \mathbf{A}_{:,k} \mathbf{A}_{:,k}^\top \mathbb{1}\{i_l = k\}$ . These matrices satisfy

$$\|Z_l\|_2 \leq \max_k \frac{\|\mathbf{A}_{:,k}\|_2^2}{rp_k} \asymp \frac{1}{r} \sum_{k=1}^n \|\mathbf{A}_{:,k}\|_2^2 = \frac{1}{\beta r} \|\mathbf{A}\|_F^2 := R$$

$$\begin{aligned} \left\| \mathbb{E} \left[ \sum_{l=1}^r Z_l Z_l^\top \right] \right\| &= \left\| r \sum_{k=1}^n \mathbb{P}\{i_l = k\} \frac{\|\mathbf{A}_{:,k}\|_2^2}{r^2 p_k^2} \mathbf{A}_{:,k} \mathbf{A}_{:,k}^\top \right\| \\ &= \frac{1}{\beta r} \|\mathbf{A}\|_F^2 \|\mathbf{A} \mathbf{A}^\top\| \\ &\leq \frac{1}{\beta r} \|\mathbf{A}\|_F^2 \|\mathbf{A}\|^2 := V \end{aligned}$$

Invoke matrix Bernstein to conclude that

$$\begin{aligned} \|\mathbf{M} - \mathbf{A} \mathbf{A}^\top\| &= \left\| \sum_{l=1}^r (Z_l - \mathbb{E}[Z_l]) \right\| \lesssim \sqrt{V \log n} + B \log n \\ &\asymp \sqrt{\frac{\log n}{\beta r}} \|\mathbf{A}\|_F \|\mathbf{A}\| \end{aligned}$$

# Matrix multiplication with one-sided information

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What if we can only use information about  $\mathbf{A}$ ?

For example, suppose  $p_k \geq \frac{\beta \|\mathbf{A}_{:,k}\|_2^2}{\|\mathbf{A}\|_F^2}$ . In this case, matrix Bernstein inequality does NOT yield sharp concentration. But we can still use Markov's inequality to get some bound

## Matrix multiplication with one-sided information

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More precisely, when  $p_k \geq \frac{\beta \|\mathbf{A}_{:,k}\|_2^2}{\|\mathbf{A}\|_F^2}$ , it follows from (1.2) that

$$\begin{aligned}\mathbb{E} \left[ \|\mathbf{M} - \mathbf{AB}\|_F^2 \right] &= \frac{1}{r} \sum_k \frac{1}{p_k} \|\mathbf{A}_{:,k}\|_2^2 \|\mathbf{B}_{k,:}\|_2^2 - \frac{1}{r} \|\mathbf{AB}\|_F^2 \\ &\leq \frac{1}{\beta r} \left( \sum_k \|\mathbf{B}_{k,:}\|_2^2 \right) \|\mathbf{A}\|_F^2 \\ &= \frac{\|\mathbf{A}\|_F^2 \|\mathbf{B}\|_F^2}{\beta r}\end{aligned}$$

Hence, Markov's inequality yields that with prob. at least  $1 - \frac{1}{\log n}$ ,

$$\|\mathbf{M} - \mathbf{AB}\|_F^2 \leq \frac{\|\mathbf{A}\|_F^2 \|\mathbf{B}\|_F^2 \log n}{\beta r} \quad (1.3)$$

## **Least squares approximation**



# Least squares (LS) problems

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Given  $\mathbf{A} \in \mathbb{R}^{n \times d}$  ( $n \gg d$ ) and  $\mathbf{b} \in \mathbb{R}^d$ , find the “best” vector s.t.  $\mathbf{Ax} \approx \mathbf{b}$ , i.e.

$$\text{minimize}_{\mathbf{x} \in \mathbb{R}^d} \quad \|\mathbf{Ax} - \mathbf{b}\|_2$$

If  $\mathbf{A}$  has full column rank, then

$$\mathbf{x}_{\text{ls}} = (\mathbf{A}^\top \mathbf{A})^{-1} \mathbf{A}^\top \mathbf{b} = \mathbf{V}_A \boldsymbol{\Sigma}_A^{-1} \mathbf{U}_A^\top \mathbf{b}$$

where  $\mathbf{A} = \mathbf{U}_A \boldsymbol{\Sigma}_A \mathbf{V}_A^\top$  is SVD of  $\mathbf{A}$ .

# Methods for solving LS problems

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**Direct methods:** computational complexity  $O(nd^2)$

- *Cholesky decomposition:* compute upper triangular matrix  $\mathbf{R}$  s.t.  $\mathbf{A}^\top \mathbf{A} = \mathbf{R}^\top \mathbf{R}$ , and solve  $\mathbf{R}^\top \mathbf{R} \mathbf{x} = \mathbf{A}^\top \mathbf{b}$
- *QR decomposition:* compute QR decomposition  $\mathbf{A} = \mathbf{Q} \mathbf{R}$  ( $\mathbf{Q}$ : orthonormal;  $\mathbf{R}$ : upper triangular), and solve  $\mathbf{R} \mathbf{x} = \mathbf{Q}^\top \mathbf{b}$

**Iterative methods:** computational complexity  $O\left(\frac{\sigma_{\max}(\mathbf{A})}{\sigma_{\min}(\mathbf{A})} \|\mathbf{A}\|_0 \log \frac{1}{\varepsilon}\right)$

- *conjugate gradient ...*

# Randomized least squares approximation

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**Basic idea:** generate sketching / sampling matrix  $\Phi$  (e.g. via random sampling, random projection), and solve instead

$$\tilde{\mathbf{x}}_{\text{ls}} = \arg \min_{\mathbf{x} \in \mathbb{R}^d} \|\Phi(\mathbf{A}\mathbf{x} - \mathbf{b})\|_2$$

**Goal:** find  $\Phi$  s.t.

$$\begin{aligned}\tilde{\mathbf{x}}_{\text{ls}} &\approx \mathbf{x}_{\text{ls}} \\ \|\mathbf{A}\tilde{\mathbf{x}}_{\text{ls}} - \mathbf{b}\|_2 &\approx \|\mathbf{A}\mathbf{x}_{\text{ls}} - \mathbf{b}\|_2\end{aligned}$$

# Which sketching matrices enable good approximation?

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We will start with two **deterministic** conditions that promise reasonably good approximation (Drineas et al '11)

# Which sketching matrices enable good approximation?

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Let  $A = U_A \Sigma_A V_A^T$  be SVD of  $A$  ...

- **Condition 1 (approximate isometry)**

$$\sigma_{\min}^2(\Phi U_A) \geq \frac{1}{\sqrt{2}} \quad (1.4)$$

- says that  $\Phi U_A$  is approximate isometry / rotation
- $1/\sqrt{2}$  can be replaced by other positive constants

# Which sketching matrices enable good approximation?

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Let  $A = U_A \Sigma_A V_A^T$  be SVD of  $A$  ...

- **Condition 2 (approximate orthogonality)**

$$\left\| U_A^T \Phi^T \Phi (Ax_{ls} - b) \right\|_2^2 \leq \frac{\varepsilon}{2} \|Ax_{ls} - b\|_2^2 \quad (1.5)$$

- says that  $\Phi U_A$  is roughly orthogonal to  $\Phi \underbrace{(Ax_{ls} - b)}_{=(U_A U_A^T - I)b}$
- even though this condition depends on  $b$ , one can find  $\Phi$  satisfying this condition without using any information about  $b$

# Can these conditions be satisfied?

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Two extreme examples

1.  $\Phi = I$ , which satisfies

$$\begin{cases} \sigma_{\min}(\Phi U_A) & = \sigma_{\min}(U_A) = 1 \\ \left\| U_A^T \Phi^T \Phi (Ax_{ls} - b) \right\|_2 & = \left\| U_A^T (I - U_A U_A^T) b \right\|_2 = 0 \end{cases}$$

- easy to construct; hard to solve subsampled LS problem

# Can these conditions be satisfied?

---

Two extreme examples

2.  $\Phi = U_A^\top$ , which satisfies

$$\begin{cases} \sigma_{\min}(\Phi U_A) & = \sigma_{\min}(\mathbf{I}) = 1 \\ \left\| U_A^\top \Phi^\top \Phi (\mathbf{A}x_{\text{ls}} - \mathbf{b}) \right\|_2 & = \left\| U_A^\top (\mathbf{I} - U_A U_A^\top) \mathbf{b} \right\|_2 = 0 \end{cases}$$

- hard to construct (i.e. compute  $U_A$ ); easy to solve subsampled LS problem



# Quality of approximation

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We'd like to assess quality of approximation w.r.t. both fitting error and estimation error

## Lemma 1.4

*Under Conditions 1-2, solution  $\tilde{\mathbf{x}}_{\text{ls}}$  to subsampled LS problem obeys*

- (i)  $\|\mathbf{A}\tilde{\mathbf{x}}_{\text{ls}} - \mathbf{b}\|_2 \leq (1 + \varepsilon)\|\mathbf{A}\mathbf{x}_{\text{ls}} - \mathbf{b}\|_2$
- (ii)  $\|\tilde{\mathbf{x}}_{\text{ls}} - \mathbf{x}_{\text{ls}}\|_2 \leq \frac{\sqrt{\varepsilon}}{\sigma_{\min}(\mathbf{A})}\|\mathbf{A}\mathbf{x}_{\text{ls}} - \mathbf{b}\|_2$

## Proof of Lemma 1.4(i)

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Subsampled LS problem can be rewritten as

$$\begin{aligned}\min_{\mathbf{x} \in \mathbb{R}^d} \|\Phi \mathbf{b} - \Phi \mathbf{A} \mathbf{x}\|_2^2 &= \min_{\Delta \in \mathbb{R}^d} \|\Phi \mathbf{b} - \Phi \mathbf{A} (\mathbf{x}_{\text{ls}} + \Delta)\|_2^2 \\ &= \min_{\Delta \in \mathbb{R}^d} \|\Phi (\mathbf{b} - \mathbf{A} \mathbf{x}_{\text{ls}}) - \Phi \mathbf{A} \Delta\|_2^2 \\ &= \min_{\mathbf{z} \in \mathbb{R}^d} \left\| \Phi (\mathbf{b} - \mathbf{A} \mathbf{x}_{\text{ls}}) - \underbrace{\Phi \mathbf{U}_A \mathbf{z}}_{=\mathbf{A}(\mathbf{x} - \mathbf{x}_{\text{ls}})} \right\|_2^2.\end{aligned}$$

Therefore, optimal solution  $\mathbf{z}_{\text{ls}}$  obeys

$$\mathbf{z}_{\text{ls}} = (\mathbf{U}_A^\top \Phi^\top \Phi \mathbf{U}_A)^{-1} (\mathbf{U}_A^\top \Phi^\top) \Phi (\mathbf{b} - \mathbf{A} \mathbf{x}_{\text{ls}}).$$

Combine Conditions 1-2 to obtain

$$\|\mathbf{z}_{\text{ls}}\|_2^2 \leq \left\| (\mathbf{U}_A^\top \Phi^\top \Phi \mathbf{U}_A)^{-1} \right\|^2 \left\| \mathbf{U}_A^\top \Phi^\top \Phi (\mathbf{b} - \mathbf{A} \mathbf{x}_{\text{ls}}) \right\|_2^2 \leq \varepsilon \|\mathbf{b} - \mathbf{A} \mathbf{x}_{\text{ls}}\|_2^2$$

## Proof of Lemma 1.4(i) (cont.)

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Previous bounds further yield

$$\begin{aligned}\|\mathbf{b} - \mathbf{A}\tilde{\mathbf{x}}_{ls}\|_2^2 &= \left\| \underbrace{\mathbf{b} - \mathbf{A}\mathbf{x}_{ls}}_{\perp U_A} + \underbrace{\mathbf{A}\mathbf{x}_{ls} - \mathbf{A}\tilde{\mathbf{x}}_{ls}}_{\in \text{range}(U_A)} \right\|_2^2 \\ &= \|\mathbf{b} - \mathbf{A}\mathbf{x}_{ls}\|_2^2 + \|\mathbf{A}\mathbf{x}_{ls} - \mathbf{A}\tilde{\mathbf{x}}_{ls}\|_2^2 \\ &= \|\mathbf{b} - \mathbf{A}\mathbf{x}_{ls}\|_2^2 + \|\mathbf{U}_A \mathbf{z}_{ls}\|_2^2 \\ &\leq \|\mathbf{b} - \mathbf{A}\mathbf{x}_{ls}\|_2^2 + \|\mathbf{z}_{ls}\|_2^2 \\ &\leq (1 + \varepsilon) \|\mathbf{b} - \mathbf{A}\mathbf{x}_{ls}\|_2^2\end{aligned}$$

Finally, we conclude proof by recognizing that  $\sqrt{1 + \varepsilon} \leq 1 + \varepsilon$ .

## Proof of Lemma 1.4(ii)

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From proof of Lemma 1.4(i), we know  $\mathbf{A}\mathbf{x}_{ls} - \mathbf{A}\tilde{\mathbf{x}}_{ls} = \mathbf{U}_A\mathbf{z}_{ls}$  and  $\|\mathbf{z}_{ls}\|_2^2 \leq \varepsilon\|\mathbf{b} - \mathbf{A}\mathbf{x}_{ls}\|_2^2$ . These reveal that

$$\begin{aligned}\|\mathbf{x}_{ls} - \tilde{\mathbf{x}}_{ls}\|_2^2 &\leq \frac{\|\mathbf{A}(\mathbf{x}_{ls} - \tilde{\mathbf{x}}_{ls})\|_2^2}{\sigma_{\min}^2(\mathbf{A})} \\ &= \frac{\|\mathbf{U}_A\mathbf{z}_{ls}\|_2^2}{\sigma_{\min}^2(\mathbf{A})} \\ &\leq \frac{\|\mathbf{z}_{ls}\|_2^2}{\sigma_{\min}^2(\mathbf{A})} \\ &\leq \frac{\varepsilon\|\mathbf{b} - \mathbf{A}\mathbf{x}_{ls}\|_2^2}{\sigma_{\min}^2(\mathbf{A})}\end{aligned}$$

## Quality of approximation (cont.)

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By making further assumption on  $\mathbf{b}$ , we can connect error bound with  $\|\mathbf{x}_{\text{ls}}\|_2$

### Lemma 1.5

Suppose  $\|\mathbf{U}_A \mathbf{U}_A^\top \mathbf{b}\|_2 \geq \gamma \|\mathbf{b}\|_2$  for some  $0 < \gamma \leq 1$ . Under Conditions 1-2, solution  $\tilde{\mathbf{x}}_{\text{ls}}$  to subsampled LS problem obeys

$$\|\mathbf{x}_{\text{ls}} - \tilde{\mathbf{x}}_{\text{ls}}\|_2 \leq \sqrt{\varepsilon} \kappa(\mathbf{A}) \sqrt{\gamma^{-2} - 1} \|\mathbf{x}_{\text{ls}}\|_2$$

where  $\kappa(\mathbf{A})$ : condition number of  $\mathbf{A}$

- $\|\mathbf{U}_A \mathbf{U}_A^\top \mathbf{b}\|_2 \geq \gamma \|\mathbf{b}\|_2$  says a nontrivial fraction of energy of  $\mathbf{b}$  lies in  $\text{range}(\mathbf{A})$

## Proof of Lemma 1.5

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Since  $\mathbf{b} - \mathbf{Ax}_{\text{ls}} = (\mathbf{I} - \mathbf{U}_A \mathbf{U}_A^\top) \mathbf{b}$ , one has

$$\begin{aligned} \|\mathbf{b} - \mathbf{Ax}_{\text{ls}}\|_2^2 &= \|(\mathbf{I} - \mathbf{U}_A \mathbf{U}_A^\top) \mathbf{b}\|_2^2 \\ &= \|\mathbf{b}\|_2^2 - \|\mathbf{U}_A \mathbf{U}_A^\top \mathbf{b}\|_2^2 \\ &\leq (\gamma^{-2} - 1) \|\mathbf{U}_A \mathbf{U}_A^\top \mathbf{b}\|_2^2 && \text{(since } \|\mathbf{U}_A \mathbf{U}_A^\top \mathbf{b}\|_2 \geq \gamma \|\mathbf{b}\|_2 \text{)} \\ &= (\gamma^{-2} - 1) \|\mathbf{Ax}_{\text{ls}}\|_2^2 && \text{(since } \mathbf{Ax}_{\text{ls}} = \mathbf{U}_A \mathbf{U}_A^\top \mathbf{b} \text{)} \\ &\leq (\gamma^{-2} - 1) \sigma_{\max}^2(\mathbf{A}) \|\mathbf{x}_{\text{ls}}\|_2^2 \end{aligned}$$

This combined with Lemma 1.4(ii) concludes proof.

# Connection with approximate matrix multiplication

Condition 1 can be guaranteed if

$$\left\| U_A^\top (\Phi^\top \Phi) U_A - \underbrace{U_A^\top U_A}_{=I} \right\| \leq 1 - \frac{1}{\sqrt{2}}$$

Condition 2 can be guaranteed if

$$\left\| U_A^\top (\Phi^\top \Phi) (Ax_{ls} - b) - \underbrace{U_A^\top (Ax_{ls} - b)}_{=U_A^\top (I - U_A U_A^\top) b = 0} \right\|_2^2 \leq \frac{\epsilon}{2} \underbrace{\|U_A\|_2^2}_{=1} \|Ax_{ls} - b\|_2^2$$

Both conditions can be viewed as approximating matrix multiplication  
(by designing  $\Phi\Phi^\top$ )

# A (slow) random projection strategy

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**Gaussian sampling:** let  $\Phi \in \mathbb{R}^{r \times n}$  be composed of i.i.d. Gaussian entries  $\mathcal{N}(0, \frac{1}{r})$

- Conditions 1-2 are satisfied with high prob. if  $r \gtrsim \frac{d \log d}{\epsilon}$   
(exercise)
- implementing Gaussian sketching is expensive (computing  $\Phi A$  takes time  $\Omega(nrd) = \Omega(nd^2 \log d)$ )



## Another random subsampling strategy

---

*Let's begin with Condition 1 and try Algorithm 1.2 with optimal sampling probabilities ...*

## Another random subsampling strategy

---

Leverage scores of  $\mathbf{A}$  are defined to be  $\|(\mathbf{U}_A)_{:,i}\|_2$  ( $1 \leq i \leq n$ )

**Nonuniform random subsampling:** set  $\Phi \in \mathbb{R}^{r \times n}$  to be a (weighted) random subsampling matrix s.t.

$$\mathbb{P} \left( \Phi_{i,:} = \frac{1}{\sqrt{r p_k}} \mathbf{e}_k^\top \right) = p_k, \quad 1 \leq k \leq n$$

with  $p_k \propto \|(\mathbf{U}_A)_{i,:}\|_2^2$

- still slow: needs to compute (exactly) leverage scores

# Fast and data-agnostic sampling

---

Can we design **data-agnostic** sketching matrix  $\Phi$  (i.e. independent of  $A, b$ ) that allows **fast** computation while satisfying Conditions 1-2?

# Subsampled randomized Hadamard transform (SRHT)

---

An SRHT matrix  $\Phi \in \mathbb{R}^{r \times n}$  is

$$\Phi = RHD$$

- $D \in \mathbb{R}^{n \times n}$ : diagonal matrix, whose entries are random  $\{\pm 1\}$
- $H \in \mathbb{R}^{n \times n}$ : Hadamard matrix (scaled by  $1/\sqrt{n}$  so it's orthonormal)
- $R \in \mathbb{R}^{r \times n}$ : uniform random subsampling

$$\mathbb{P}\left(R_{i,:} = \sqrt{\frac{n}{r}} \mathbf{e}_k^\top\right) = \frac{1}{n}, \quad 1 \leq k \leq n$$

# Subsampled randomized Hadamard transform

---

## Key idea of SRHT:

- use  $HD$  to “uniformize” leverage scores (so that  $\{\|(HDU_A)_{i,:}\|_2\}$  are more-or-less identical)
- subsample rank-one components **uniformly** at random

# Uniformization of leverage scores

## Lemma 1.6

For any fixed matrix  $U \in \mathbb{R}^{n \times d}$ , one has

$$\max_{1 \leq i \leq n} \|(\mathbf{H}DU)_{i,:}\|_2 \lesssim \frac{\log n}{\sqrt{n}} \|U\|_F$$

with prob. exceeding  $1 - O(n^{-9})$

- $\mathbf{H}D$  preconditions  $U$  with high prob.; more precisely,

$$\frac{\|(\mathbf{H}DU)_{i,:}\|_2^2}{\sum_{l=1}^n \|(\mathbf{H}DU)_{l,:}\|_2^2} = \frac{\|(\mathbf{H}DU)_{i,:}\|_2^2}{\|U\|_F^2} \lesssim \frac{\log^2 n}{n} \quad (1.6)$$

## Proof of Lemma 1.6

---

For any fixed matrix  $\mathbf{U} \in \mathbb{R}^{n \times d}$ , one has

$$(\mathbf{H}\mathbf{D}\mathbf{U})_{i,:} = \sum_{j=1}^n \underbrace{h_{i,j} D_{j,j}}_{\text{random on } \{\pm \frac{1}{\sqrt{n}}\}} \mathbf{U}_{j,:},$$

which clearly satisfies  $\mathbb{E}[(\mathbf{H}\mathbf{D}\mathbf{U})_{i,:}] = \mathbf{0}$ . In addition,

$$V := \mathbb{E} \left[ \sum_{j=1}^n \|h_{i,j} D_{j,j} \mathbf{U}_{j,:}\|_2^2 \right] = \frac{1}{n} \sum_{j=1}^n \|\mathbf{U}_{j,:}\|_2^2 = \frac{1}{n} \|\mathbf{U}\|_{\text{F}}^2$$

$$B := \max_j \|h_{i,j} D_{j,j} \mathbf{U}_{j,:}\|_2 = \frac{1}{\sqrt{n}} \max_j \|\mathbf{U}_{j,:}\|_2 \leq \frac{1}{\sqrt{n}} \|\mathbf{U}\|_{\text{F}}$$

Invoke matrix Bernstein to demonstrate that with prob.  $1 - O(n^{-10})$ ,

$$\|(\mathbf{H}\mathbf{D}\mathbf{U})_{i,:}\|_2 \lesssim \sqrt{V \log n} + B \log n \lesssim \frac{\log n}{\sqrt{n}} \|\mathbf{U}\|_{\text{F}}$$

# Theoretical guarantees for SRHT

---

When uniform subsampling is adopted, one has  $p_k = 1/n$ . In view of Lemma 1.6,

$$p_k \geq \beta \frac{\|(\mathbf{H}\mathbf{D}\mathbf{U}_A)_{i,:}\|_2^2}{\sum_{l=1}^n \|(\mathbf{H}\mathbf{D}\mathbf{U}_A)_{l,:}\|_2^2}$$

with  $\beta \asymp \log^{-2} n$ . Apply Theorem 1.3 to yield

$$\begin{aligned} \left\| \mathbf{U}_A^\top \Phi^\top \Phi \mathbf{U}_A - \mathbf{I} \right\| &= \left\| \mathbf{U}_A^\top \Phi^\top \Phi \mathbf{U}_A - \mathbf{U}_A^\top \mathbf{U}_A \right\| \\ &= \left\| (\mathbf{U}_A^\top \mathbf{D}^\top \mathbf{H}^\top) \mathbf{R}^\top \mathbf{R} (\mathbf{H}\mathbf{D}\mathbf{U}_A) - (\mathbf{U}_A^\top \mathbf{D}^\top \mathbf{H}^\top) (\mathbf{H}\mathbf{D}\mathbf{U}_A) \right\| \\ &\leq 1/2 \end{aligned}$$

when  $r \gtrsim \frac{\|\mathbf{H}\mathbf{D}\mathbf{U}_A\|_F^2 \log n}{\|\mathbf{H}\mathbf{D}\mathbf{U}_A\|_2^2 \beta} \asymp d \log^3 n$ . This establishes Condition 1



# Theoretical guarantees for SRHT

---

Similarly, Condition 2 is satisfied with high prob. if  $r \gtrsim \frac{d \log^3 n}{\varepsilon}$   
(exercise)

# Back to least squares approximation

---

Preceding analysis suggests following algorithm

---

**Algorithm 1.3** Randomized LS approximation (uniform sampling)

---

- 1: Pick  $r \gtrsim \frac{d \log^3 n}{\varepsilon}$ , and generate  $\mathbf{R} \in \mathbb{R}^{r \times n}$ ,  $\mathbf{H} \in \mathbb{R}^{n \times n}$  and  $\mathbf{D} \in \mathbb{R}^{n \times n}$  (as described before)
  - 2: **return**  $\tilde{\mathbf{x}} = (\mathbf{R}\mathbf{H}\mathbf{D}\mathbf{A})^\dagger \mathbf{R}\mathbf{H}\mathbf{D}\mathbf{b}$
- 

- computational complexity:

$$O\left( \underbrace{nd \log \frac{n}{\varepsilon}}_{\text{compute } \mathbf{HDA}} + \underbrace{\frac{d^3 \log^3 n}{\varepsilon}}_{\text{solve subsampled LS } (rd^2)} \right)$$

# An alternative approach: nonuniform sampling

---

Key idea of Algorithm 1.3 is to uniformize leverage scores followed by uniform sampling

Alternatively, one can also start by estimating leverage scores, and then apply **nonuniform sampling** accordingly

# Fast approximation of leverage scores

---

**Key idea:** apply SRHT (or other fast Johnson-Lindenstrass transform) in appropriate places

$$\begin{aligned}\|U_{i,:}\|_2^2 &= \|e_i^\top U\|_2^2 = \|e_i^\top U U^\top\|_2^2 \\ &= \|e_i^\top \mathbf{A} \mathbf{A}^\dagger\|_2^2 \\ &= \|e_i^\top \mathbf{A} \mathbf{A}^\dagger \Phi_1^\top\|_2^2\end{aligned}$$

where  $\Phi_1 \in \mathbb{R}^{r_1 \times n}$  is SRHT matrix

**Issue:**  $\mathbf{A} \mathbf{A}^\dagger$  is expensive to compute; can we compute  $\mathbf{A} \mathbf{A}^\dagger \Phi_1^\top$  in a fast manner?

## Aside: pseudo inverse

---

Let  $\Phi \in \mathbb{R}^{r \times n}$  be SRHT matrix with sufficiently large  $r \gg \frac{d \text{poly} \log n}{\varepsilon^2}$ .  
With high prob., one has (check Mahoney's lecture notes)

$$\|(\Phi U_A)^\dagger - (\Phi U_A)^\top\| \leq \varepsilon$$

$$\text{and } (\Phi A)^\dagger = V_A \Sigma_A^{-1} (\Phi U_A)^\dagger$$

These mean

$$\begin{aligned} A(\Phi A)^\dagger &= U_A \Sigma_A V_A^\top V_A \Sigma_A^{-1} (\Phi U_A)^\dagger \approx U_A \Sigma_A V_A^\top V_A \Sigma_A^{-1} (\Phi U_A)^\top \\ &= U_A U_A^\top \Phi^\top = A A^\dagger \Phi \end{aligned}$$

# Fast approximation of leverage scores

---

**Continuing our key idea:** apply SRHT (or other fast Johnson-Lindenstrass transform) in appropriate places

$$\begin{aligned}\|U_{i,:}\|_2^2 &\approx \|e_i^\top \mathbf{A}(\Phi_1 \mathbf{A})^\dagger\|_2^2 \\ &\approx \|e_i^\top \mathbf{A}(\Phi_1 \mathbf{A})^\dagger \Phi_2\|_2^2\end{aligned}$$

where  $\Phi_1 \in \mathbb{R}^{r_1 \times n}$  and  $\Phi_2 \in \mathbb{R}^{r_1 \times r_2}$  ( $r_2 \asymp \text{poly log } n$ ) are both SRHT matrices

# Fast approximation of leverage scores

---

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## Algorithm 1.4 Leverage scores approximation

---

- 1: Pick  $r_1 \gtrsim \frac{d \log^3 n}{\varepsilon}$  and  $r_2 \asymp \text{poly log } n$
  - 2: Compute  $\Phi_1 \mathbf{A} \in \mathbb{R}^{r_1 \times d}$  and its QR decomposition, and let  $\mathbf{R}_{\Phi_1 \mathbf{A}}$  be the “R” matrix from QR
  - 3: Construct  $\Psi = \mathbf{A} \mathbf{R}_{\Phi_1 \mathbf{A}}^{-1} \Phi_2$
  - 4: **return**  $\ell_i = \|\Psi_{i,:}\|_2$
- 

- computational complexity:  $O\left(\frac{nd \text{poly log } n}{\varepsilon^2} + \frac{d^3 \text{poly log } n}{\varepsilon^2}\right)$

# Least squares approximation (nonuniform sampling)

---

---

## Algorithm 1.5 Randomized LS approximation (nonuniform sampling)

---

- 1: Run Algorithm 1.4 to compute approximate leverage scores  $\{\ell_k\}$ , and set  $p_k \propto \ell_k^2$
  - 2: Randomly sample  $r \gtrsim \frac{d \text{poly} \log n}{\varepsilon}$  rows of  $\mathbf{A}$  and elements of  $\mathbf{b}$  using  $\{p_k\}$  as sampling probabilities, rescaling each by  $1/\sqrt{rp_k}$ . Let  $\Phi \mathbf{A}$  and  $\Phi \mathbf{b}$  be the subsampled matrix and vector
  - 3: **return**  $\tilde{\mathbf{x}}_{\text{ls}} = \arg \min_{\mathbf{x} \in \mathbb{R}^d} \|\Phi \mathbf{A} \mathbf{x} - \Phi \mathbf{b}\|_2$
- 

informally, Algorithm 1.5 returns a reasonably good solution with prob.  $1 - O(1/\log n)$



## **Low-rank matrix approximation**

# Low-rank matrix approximation

---

**Question:** given a matrix  $A \in \mathbb{R}^{n \times n}$ , how to find a rank- $k$  matrix that well approximates  $A$

- One can compute SVD of  $A = U\Sigma V^\top$ , then return

$$A_k = U_k U_k^\top A$$

where  $U_k$  consists of top- $k$  singular vectors

- In general, takes time  $O(n^3)$ , or  $O(kn^2)$  (by power methods)
- Can we find faster algorithms if we only want “good approximation”?

# Randomized low-rank matrix approximation

---

**Strategy:** find a matrix  $C$  (via, e.g., subsampling columns of  $A$ ), and return

$$\underbrace{CC^\dagger A}$$

project  $A$  onto column space of  $C$

**Question:** how well can  $CC^\dagger A$  approximate  $A$ ?

# A simple paradigm

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## Algorithm 1.6

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- 1: **input:** data matrix  $A \in \mathbb{R}^{n \times n}$ , subsampled matrix  $C \in \mathbb{R}^{n \times r}$
  - 2: **return**  $H_k$  as top- $k$  left singular vectors of  $C$
- 

- As we will see, quality of approximation depends on size of

$$\underbrace{AA^T - CC^T}$$

connection with matrix multiplication

# Quality of approximation (Frobenius norm)

---

One can also connect spectral-norm error with product of matrices

## Lemma 1.7

*The output of Algorithm 1.6 satisfies*

$$\|\mathbf{A} - \mathbf{H}_k \mathbf{H}_k^\top \mathbf{A}\|_F^2 \leq \|\mathbf{A} - \mathbf{U}_k \mathbf{U}_k^\top \mathbf{A}\|_F^2 + 2\sqrt{k} \|\mathbf{A} \mathbf{A}^\top - \mathbf{C} \mathbf{C}^\top\|_F$$

where  $\mathbf{U}_k \in \mathbb{R}^{n \times k}$  contains top- $k$  left singular vectors of  $\mathbf{A}$

- This holds for any  $\mathbf{C}$
- Approximation error depends on the error in approximating product of two matrices

## Proof of Lemma 1.7

---

To begin with, since  $\mathbf{H}_k$  is orthonormal, one has

$$\|\mathbf{A} - \mathbf{H}_k \mathbf{H}_k^\top \mathbf{A}\|_{\text{F}}^2 = \|\mathbf{A}\|_{\text{F}}^2 - \|\mathbf{H}_k^\top \mathbf{A}\|_{\text{F}}^2$$

Next, letting  $\mathbf{h}_i = (\mathbf{H}_k)_{:,i}$  yields

$$\begin{aligned} \left| \|\mathbf{H}_k^\top \mathbf{A}\|_{\text{F}}^2 - \sum_{i=1}^k \sigma_i^2(\mathbf{C}) \right| &= \left| \sum_{i=1}^k \|\mathbf{A}^\top \mathbf{h}_i\|_2^2 - \sum_{i=1}^k \|\mathbf{C} \mathbf{h}_i\|_2^2 \right| \\ &= \left| \sum_{i=1}^k \langle \mathbf{h}_i \mathbf{h}_i^\top, \mathbf{A} \mathbf{A}^\top - \mathbf{C} \mathbf{C}^\top \rangle \right| \\ &= \left| \langle \mathbf{H}_k \mathbf{H}_k^\top, \mathbf{A} \mathbf{A}^\top - \mathbf{C} \mathbf{C}^\top \rangle \right| \\ &\leq \|\mathbf{H}_k \mathbf{H}_k^\top\|_{\text{F}} \|\mathbf{A} \mathbf{A}^\top - \mathbf{C} \mathbf{C}^\top\|_{\text{F}} \\ &\leq \sqrt{k} \|\mathbf{A} \mathbf{A}^\top - \mathbf{C} \mathbf{C}^\top\|_{\text{F}} \end{aligned}$$

## Proof of Lemma 1.7

---

In addition,

$$\begin{aligned} \left| \sum_{i=1}^k \sigma_i^2(\mathbf{C}) - \sum_{i=1}^k \sigma_i^2(\mathbf{A}) \right| &= \left| \sum_{i=1}^k \left\{ \sigma_i(\mathbf{C}\mathbf{C}^\top) - \sigma_i(\mathbf{A}\mathbf{A}^\top) \right\} \right| \\ &\leq \sqrt{k} \sqrt{\sum_{i=1}^n \left\{ \sigma_i(\mathbf{C}\mathbf{C}^\top) - \sigma_i(\mathbf{A}\mathbf{A}^\top) \right\}^2} \quad (\text{Cauchy-Schwarz}) \\ &\leq \sqrt{k} \|\mathbf{C}\mathbf{C}^\top - \mathbf{A}\mathbf{A}^\top\|_{\text{F}} \quad (\text{Wielandt-Hoffman inequality}) \end{aligned}$$

Finally, one has  $\|\mathbf{A} - \mathbf{U}_k \mathbf{U}_k^\top \mathbf{A}\|_{\text{F}}^2 = \|\mathbf{A}\|_{\text{F}}^2 - \sum_{i=1}^k \sigma_i^2(\mathbf{A})$ .

Combining above results establishes the claim

# Quality of approximation (spectral norm)

---

## Lemma 1.8

*The output of Algorithm 1.6 satisfies*

$$\|A - H_k H_k^\top A\|^2 \leq \|A - U_k U_k^\top A\|^2 + 2\|AA^\top - CC^\top\|$$

*where  $U_k \in \mathbb{R}^{n \times k}$  contains top- $k$  left singular vectors of  $A$*



## Proof of Lemma 1.8

---

First of all,

$$\begin{aligned}\|\mathbf{A} - \mathbf{H}_k \mathbf{H}_k^\top \mathbf{A}\| &= \max_{\mathbf{x}: \|\mathbf{x}\|_2=1} \|\mathbf{x}^\top (\mathbf{I} - \mathbf{H}_k \mathbf{H}_k^\top) \mathbf{A}\|_2 \\ &= \max_{\mathbf{x}: \|\mathbf{x}\|_2=1, \mathbf{x} \perp \mathbf{H}_k} \|\mathbf{x}^\top \mathbf{A}\|_2\end{aligned}$$

Additionally, for any  $\mathbf{x} \perp \mathbf{H}_k$ ,

$$\begin{aligned}\|\mathbf{x}^\top \mathbf{A}\|_2^2 &= \left| \mathbf{x}^\top \mathbf{C} \mathbf{C}^\top \mathbf{x} + \mathbf{x}^\top (\mathbf{A} \mathbf{A}^\top - \mathbf{C} \mathbf{C}^\top) \mathbf{x} \right| \\ &\leq \left| \mathbf{x}^\top \mathbf{C} \mathbf{C}^\top \mathbf{x} \right| + \left| \mathbf{x}^\top (\mathbf{A} \mathbf{A}^\top - \mathbf{C} \mathbf{C}^\top) \mathbf{x} \right| \\ &\leq \sigma_{k+1}(\mathbf{C} \mathbf{C}^\top) + \|\mathbf{A} \mathbf{A}^\top - \mathbf{C} \mathbf{C}^\top\| \\ &\leq \sigma_{k+1}(\mathbf{A} \mathbf{A}^\top) + 2\|\mathbf{A} \mathbf{A}^\top - \mathbf{C} \mathbf{C}^\top\| \\ &= \|\mathbf{A} - \mathbf{U}_k \mathbf{U}_k^\top \mathbf{A}\|^2 + 2\|\mathbf{A} \mathbf{A}^\top - \mathbf{C} \mathbf{C}^\top\|.\end{aligned}$$

This concludes the proof.

# Back to low-rank matrix approximation

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To ensure  $\mathbf{A}\mathbf{A}^\top - \mathbf{C}\mathbf{C}^\top$  is small, we can do random subsampling / projection as before. For example:

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## Algorithm 1.7

---

- 1: **for**  $l = 1, \dots, r$  **do**
  - 2:   Pick  $i_l \in \{1, \dots, n\}$  i.i.d. with prob.  $\mathbb{P}\{i_l = k\} = p_k$
  - 3:   Set  $\mathbf{C}_{:,l} = \frac{1}{\sqrt{r p_{i_l}}} \mathbf{A}_{:,l}$
  - 4: **return**  $\mathbf{H}_k$  as top- $k$  left singular vectors of  $\mathbf{C}$
-

# Back to low-rank matrix approximation

---

Invoke Theorems 1.2 and 1.3 to see that with high prob.:

- If  $r \gtrsim \frac{k \log n}{\beta \varepsilon^2}$ , then

$$\|\mathbf{A} - \mathbf{H}_k \mathbf{H}_k^\top \mathbf{A}\|_F^2 \leq \|\mathbf{A} - \mathbf{U}_k \mathbf{U}_k^\top \mathbf{A}\|_F^2 + \varepsilon \|\mathbf{A}\|_F^2 \quad (1.7)$$

- If  $r \gtrsim \frac{\|\mathbf{A}\|_F^2 \log n}{\|\mathbf{A}\|^2 \beta \varepsilon^2}$ , then

$$\|\mathbf{A} - \mathbf{H}_k \mathbf{H}_k^\top \mathbf{A}\|^2 \leq \|\mathbf{A} - \mathbf{U}_k \mathbf{U}_k^\top \mathbf{A}\|^2 + \varepsilon \|\mathbf{A}\|^2 \quad (1.8)$$

# An improved multi-pass algorithm

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## Algorithm 1.8 Multi-pass randomized SVD

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- 1:  $\mathcal{S} = \{\}$
  - 2: **for**  $l = 1, \dots, t$  **do**
  - 3:    $\mathbf{E}_l = \mathbf{A} - \mathbf{A}_{\mathcal{S}} \mathbf{A}_{\mathcal{S}}^\dagger \mathbf{A}$
  - 4:   Set  $p_k \geq \frac{\beta \|(\mathbf{E}_l)_{:,k}\|_2^2}{\|\mathbf{E}_l\|_{\text{F}}^2}$ ,  $1 \leq k \leq n$
  - 5:   Randomly select  $r$  column indices with sampling prob.  $\{p_k\}$  and append to  $\mathcal{S}$
  - 6: **return**  $\mathbf{C} = \mathbf{A}_{\mathcal{S}}$
-

# An improved multi-pass algorithm

---

## Theorem 1.9

Suppose  $r \gtrsim \frac{k \log n}{\beta \varepsilon^2}$ . With high prob.,

$$\|\mathbf{A} - \mathbf{C}\mathbf{C}^\dagger \mathbf{A}\|_{\text{F}}^2 \leq \frac{1}{1 - \varepsilon} \|\mathbf{A} - \mathbf{U}_k \mathbf{U}_k^\top\|_{\text{F}}^2 + \varepsilon^t \|\mathbf{A}\|_{\text{F}}^2$$

## Proof of Theorem 1.9

---

We will prove it by induction. Clearly, the claim holds for  $t = 1$  (according to (1.7)). Assume

$$\left\| \underbrace{\mathbf{A} - \mathbf{C}^{t-1}(\mathbf{C}^{t-1})^\dagger \mathbf{A}}_{:=\mathbf{E}_t} \right\|_{\mathbb{F}}^2 \leq \frac{1}{1-\varepsilon} \|\mathbf{A} - \mathbf{U}_k \mathbf{U}_k^\top \mathbf{A}\|_{\mathbb{F}}^2 + \varepsilon^{t-1} \|\mathbf{A}\|_{\mathbb{F}}^2,$$

and let  $\mathbf{Z}$  be the matrix of the columns of  $\mathbf{E}_t$  included in the sample. In view of (1.7),

$$\left\| \mathbf{E}_t - \mathbf{Z}\mathbf{Z}^\dagger \mathbf{E}_t \right\|_{\mathbb{F}}^2 \leq \|\mathbf{E}_t - (\mathbf{E}_t)_k\|_{\mathbb{F}}^2 + \varepsilon \|\mathbf{E}_t\|_{\mathbb{F}}^2,$$

with  $(\mathbf{E}_t)_k$  the best rank- $k$  approximation of  $\mathbf{E}_t$ . Combining the above two inequalities yields

$$\begin{aligned} \left\| \mathbf{E}_t - \mathbf{Z}\mathbf{Z}^\dagger \mathbf{E}_t \right\|_{\mathbb{F}}^2 &\leq \|\mathbf{E}_t - (\mathbf{E}_t)_k\|_{\mathbb{F}}^2 \\ &+ \frac{\varepsilon}{1-\varepsilon} \|\mathbf{A} - \mathbf{U}_k \mathbf{U}_k^\top \mathbf{A}\|_{\mathbb{F}}^2 + \varepsilon^t \|\mathbf{A}\|_{\mathbb{F}}^2 \end{aligned} \quad (1.9)$$

## Proof of Theorem 1.9 (cont.)

---

If we can show that

$$\mathbf{E}_t - \mathbf{Z}\mathbf{Z}^\dagger\mathbf{E}_t = \mathbf{A} - \mathbf{C}^t(\mathbf{C}^t)^\dagger\mathbf{A} \quad (1.10)$$

$$\|\mathbf{E}_t - (\mathbf{E}_t)_k\|_{\mathbb{F}}^2 \leq \|\mathbf{A} - \mathbf{A}_k\|_{\mathbb{F}}^2 \quad (1.11)$$

then substitution into (1.9) yields

$$\begin{aligned} \left\| \mathbf{A} - \mathbf{C}^t(\mathbf{C}^t)^\dagger\mathbf{A} \right\|_{\mathbb{F}}^2 &\leq \|\mathbf{A} - \mathbf{A}_k\|_{\mathbb{F}}^2 + \frac{\varepsilon}{1-\varepsilon} \|\mathbf{A} - \mathbf{A}_k\|_{\mathbb{F}}^2 + \varepsilon^t \|\mathbf{A} - \mathbf{A}_k\|_{\mathbb{F}}^2 \\ &= \frac{1}{1-\varepsilon} \|\mathbf{A} - \mathbf{A}_k\|_{\mathbb{F}}^2 + \varepsilon^t \|\mathbf{A} - \mathbf{A}_k\|_{\mathbb{F}}^2 \end{aligned}$$

We can then use induction to finish proof

## Proof of Theorem 1.9 (cont.)

---

It remains to justify (1.10) and (1.11).

To begin with, (1.10) follows from the definition of  $\mathbf{E}_t$  and the fact  $\mathbf{Z}\mathbf{Z}^\dagger\mathbf{C}^{t-1}(\mathbf{C}^{t-1})^\dagger = \mathbf{0}$ , which gives

$$\mathbf{C}^t(\mathbf{C}^t)^\dagger = \mathbf{C}^{t-1}(\mathbf{C}^{t-1})^\dagger + \mathbf{Z}\mathbf{Z}^\dagger$$



## Proof of Theorem 1.9 (cont.)

---

To show (1.11), note that  $(\mathbf{E}_t)_k$  is best rank- $k$  approximation of  $\mathbf{E}_t$ . This gives

$$\begin{aligned}\|\mathbf{E}_t - (\mathbf{E}_t)_k\|_{\mathbb{F}}^2 &= \|(\mathbf{I} - \mathbf{C}^{t-1}(\mathbf{C}^{t-1})^\dagger) \mathbf{A} - ((\mathbf{I} - \mathbf{C}^{t-1}(\mathbf{C}^{t-1})^\dagger) \mathbf{A})_k\|_{\mathbb{F}}^2 \\ &\leq \|(\mathbf{I} - \mathbf{C}^{t-1}(\mathbf{C}^{t-1})^\dagger) \mathbf{A} - (\mathbf{I} - \mathbf{C}^{t-1}(\mathbf{C}^{t-1})^\dagger) \mathbf{A}_k\|_{\mathbb{F}}^2 \\ &\text{(since } (\mathbf{I} - \mathbf{C}^{t-1}(\mathbf{C}^{t-1})^\dagger) \mathbf{A}_k \text{ is rank-}k\text{)} \\ &= \|(\mathbf{I} - \mathbf{C}^{t-1}(\mathbf{C}^{t-1})^\dagger) (\mathbf{A} - \mathbf{A}_k)\|_{\mathbb{F}}^2 \\ &\leq \|\mathbf{A} - \mathbf{A}_k\|_{\mathbb{F}}^2,\end{aligned}$$

where  $\mathbf{A}_k$  is best rank- $k$  approximation of  $\mathbf{A}$ . Substitution into (1.9) establishes the claim for  $t$

# Multiplicative error bounds

---

So far, our results read

$$\|\mathbf{A} - \mathbf{C}\mathbf{C}^\dagger \mathbf{A}\|_{\mathbf{F}}^2 \leq \|\mathbf{A} - \mathbf{A}_k\|_{\mathbf{F}}^2 + \text{additive error}$$

$$\|\mathbf{A} - \mathbf{C}\mathbf{C}^\dagger \mathbf{A}\|^2 \leq \|\mathbf{A} - \mathbf{A}_k\|^2 + \text{additive error}$$

In some cases, one might prefer multiplicative error guarantees, e.g.

$$\|\mathbf{A} - \mathbf{C}\mathbf{C}^\dagger \mathbf{A}\|_{\mathbf{F}} \leq (1 + \varepsilon) \|\mathbf{A} - \mathbf{A}_k\|_{\mathbf{F}}$$

## Two types of matrix decompositions

---

- *CX decomposition*: let  $\mathbf{C} \in \mathbb{R}^{n \times r}$  consist of  $r$  columns of  $\mathbf{A}$ , and return

$$\hat{\mathbf{A}} = \mathbf{C}\mathbf{X}$$

for some matrix  $\mathbf{X} \in \mathbb{R}^{r \times n}$

- *CUR decomposition*: let  $\mathbf{C} \in \mathbb{R}^{n \times r}$  (resp.  $\mathbf{R} \in \mathbb{R}^{r \times n}$ ) consist of  $r$  columns (resp. rows) of  $\mathbf{A}$ , and return

$$\hat{\mathbf{A}} = \mathbf{C}\mathbf{U}\mathbf{R}$$

for some matrix  $\mathbf{U} \in \mathbb{R}^{r \times r}$

# Generalized least squares problem

---

$$\text{minimize}_{\mathbf{X}} \quad \|\mathbf{B} - \mathbf{A}\mathbf{X}\|_{\text{F}}^2$$

where  $\mathbf{X}$  is matrix (rather than vector)

- generalization of over-determined  $\ell_2$  regression
- optimal solution:  $\mathbf{X}^{\text{ls}} = \mathbf{A}^\dagger \mathbf{B}$
- if  $\text{rank}(\mathbf{A}) \leq k$ , then  $\mathbf{X}^{\text{ls}} = \mathbf{A}_k^\dagger \mathbf{B}$

# Generalized least squares approximation

---

**Randomized algorithm:** construct an optimally weighted subsampling matrix  $\Phi \in \mathbb{R}^{r \times n}$  with  $r \gtrsim \frac{k^2}{\epsilon^2}$  and compute

$$\tilde{\mathbf{X}}^{\text{ls}} = (\Phi \mathbf{A})^\dagger \Phi \mathbf{B}$$

Then informally, with high probability,

$$\begin{aligned} \|\mathbf{B} - \mathbf{A}\tilde{\mathbf{X}}^{\text{ls}}\|_{\text{F}} &\leq (1 + \epsilon) \left\{ \min_{\mathbf{X}} \|\mathbf{B} - \mathbf{A}\mathbf{X}\|_{\text{F}} \right\} \\ \|\mathbf{X}^{\text{ls}} - \tilde{\mathbf{X}}^{\text{ls}}\|_{\text{F}} &\leq \frac{\epsilon}{\sigma_{\min}(\mathbf{A}_k)} \left\{ \min_{\mathbf{X}} \|\mathbf{B} - \mathbf{A}\mathbf{X}\|_{\text{F}} \right\} \end{aligned}$$

# Randomized algorithm for CX decomposition

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**Algorithm 1.9** Randomized algorithm for constructing CX matrix decompositions

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- 1: Compute / approximate sampling probabilities  $\{p_i\}_{i=1}^n$ , where  $p_i = \frac{1}{k} \|(\mathbf{U}_{A,k})_{:,i}\|_2^2$
  - 2: Use sampling probabilities  $\{p_i\}$  to construct a rescaled random sampling matrix  $\Phi$
  - 3: Construct  $\mathbf{C} = \mathbf{A}\Phi^\top$
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# Theoretical guarantees

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## Theorem 1.10

Suppose  $r \gtrsim \frac{k \log k}{\varepsilon^2}$ , then Algorithm 1.9 yields

$$\|\mathbf{A} - \mathbf{C}\mathbf{C}^\dagger\mathbf{A}\|_{\text{F}} \leq (1 + \varepsilon)\|\mathbf{A} - \mathbf{A}_k\|_{\text{F}}$$

# Proof of Theorem 1.10

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$$\begin{aligned} & \| \mathbf{A} - \underbrace{\mathbf{C}\mathbf{C}^\dagger}_{:= \mathbf{X}^{\text{ls}}} \mathbf{A} \|_{\text{F}} \\ &= \| \mathbf{A} - (\mathbf{A}\Phi^\top)(\mathbf{A}\Phi^\top)^\dagger \mathbf{A} \|_{\text{F}} \\ &\leq \| \mathbf{A} - (\mathbf{A}\Phi^\top)(\mathbf{P}_{A_k}\mathbf{A}\Phi^\top)^\dagger \mathbf{P}_{A_k}\mathbf{A} \|_{\text{F}} \quad (\mathbf{P}_{A_k} := \mathbf{U}_k\mathbf{U}_k^\top) \\ &\quad \text{since } \mathbf{X}^{\text{ls}} := \mathbf{C}^\dagger \mathbf{A} \text{ minimizes } \| \mathbf{A} - \mathbf{C}\mathbf{X} \|_{\text{F}} \\ &= \| \mathbf{A} - (\mathbf{A}\Phi^\top)(\mathbf{A}_k\Phi^\top)^\dagger \mathbf{A}_k \|_{\text{F}} \\ &\leq (1 + \varepsilon) \| \mathbf{A} - \mathbf{A}\mathbf{A}_k^\dagger \mathbf{A}_k \|_{\text{F}} \\ &= (1 + \varepsilon) \| \mathbf{A} - \mathbf{A}_k \|_{\text{F}} \end{aligned}$$