ELE 520: Mathematics of Data Science

## **Robust Principal Component Analysis**



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Suppose we are given a matrix



**Question:** can we hope to recover both L and S from M?

# Principal component analysis (PCA)

- N samples  $oldsymbol{X} = [oldsymbol{x}_1, oldsymbol{x}_2, \dots, oldsymbol{x}_N] \in \mathbb{R}^{n imes N}$  that are centered
- PCA: seeks r directions that explain most variance of data

minimize<sub>*L*:rank(*L*)=
$$r$$
  $\|X - L\|_{\rm F}$</sub> 

 $\circ$  best rank-r approximation of  $oldsymbol{X}$ 



# Sensitivity to corruptions / outliers

What if some samples are corrupted (e.g. due to sensor errors / attacks)?



Classical PCA fails even with a few outliers

### Video surveillance

Separation of background (low-rank) and foreground (sparse)



Candès, Li, Ma, Wright '11

# Graph clustering / community recovery

- n nodes, 2 (or more) clusters
- A friendship graph  $\mathcal{G}$ : for any pair (i, j),

$$M_{i,j} = \begin{cases} 1, & \text{if } (i,j) \in \mathcal{G} \\ 0, & \text{else} \end{cases}$$

- Edge density within clusters > edge density across clusters
- Goal: recover cluster structure



# Graph clustering / community recovery



• An equivalent goal: recover the ground truth matrix

$$L_{i,j} = \begin{cases} 1, & \text{if } i \text{ and } j \text{ are in the same community} \\ 0, & \text{else} \end{cases}$$

• Clustering  $\iff$  robust PCA

#### Fact 14.1

Consider a Gaussian vector  $\boldsymbol{x} \sim \mathcal{N}(\boldsymbol{0}, \boldsymbol{\Sigma})$ . For any u and v,

 $x_u \perp \!\!\!\perp x_v \mid \boldsymbol{x}_{\mathcal{V} \setminus \{u,v\}}$ 

iff  $\Theta_{u,v} = 0$ , where  $\Theta = \Sigma^{-1}$  is the inverse covariance matrix

#### conditional independence $\iff$ sparsity



#### The inverse covariance matrix $\Theta$ is often sparse

## Graphical models with latent factors

What if one only observes a subset of variables?

 $egin{array}{c|c} m{x}_{
m o} & (observed variables) \ m{x}_{
m h} & (hidden variables) \end{array}$ 



The covariance and precision matrices can be partitioned as

$$\boldsymbol{\Sigma} = \begin{bmatrix} \overbrace{\boldsymbol{\Sigma}_{o}}^{\text{observed part}} & \boldsymbol{\Sigma}_{o,h} \\ \boldsymbol{\Sigma}_{o,h}^\top & \boldsymbol{\Sigma}_{h} \end{bmatrix} = \begin{bmatrix} \boldsymbol{\Theta}_{o} & \boldsymbol{\Theta}_{o,h} \\ \boldsymbol{\Theta}_{o,h}^\top & \boldsymbol{\Theta}_{h} \end{bmatrix}^{-1}$$

## Graphical models with latent factors

What if one only observes a subset of variables?  $x_5$  $x_1$  $x_4$  $egin{array}{c|c} m{x}_{
m o} & (observed variables) \ m{x}_{
m h} & (hidden variables) \end{array}$  $x_2$  $x_8$  $x_6$  $x_2$  $x_7$  $\boldsymbol{x}_{0} = [x_{1}, \cdots, x_{6}]^{\top}, \boldsymbol{x}_{b} = [x_{7}, x_{8}]^{\top}$  $\underbrace{\mathbf{\Theta}_{\mathrm{o},\mathrm{h}}\mathbf{\Theta}_{\mathrm{h}}^{-1}\mathbf{\Theta}_{\mathrm{h},\mathrm{o}}}_{\mathrm{h}}$ low-rank if # latent vars is small sparse observer

sparse + low-rank decomposition

Identifiability issue: a matrix might be simultaneously low-rank and sparse!



Nonzero entries of sparse component need to be spread out

— This lecture: assume locations of the nonzero entries are random

Identifiability issue: a matrix might be simultaneously low-rank and sparse!



The low-rank component needs to be incoherent

### Definition 14.2

Coherence parameter  $\mu_1$  of  $M = U\Sigma V^{\top}$  is the smallest quantity s.t.

$$\max_{i} \|\boldsymbol{U}^{\top}\boldsymbol{e}_{i}\|_{2}^{2} \leq \frac{\mu_{1}r}{n} \quad \text{and} \quad \max_{i} \|\boldsymbol{V}^{\top}\boldsymbol{e}_{i}\|_{2}^{2} \leq \frac{\mu_{1}r}{n}$$



#### **Definition 14.3 (Joint coherence)**

Joint coherence parameter  $\mu_2$  of  $M = U \Sigma V^{\top}$  is the smallest quantity s.t.

 $\|\boldsymbol{U}\boldsymbol{V}^{\top}\|_{\infty} \leq \sqrt{\frac{\mu_2 r}{n^2}}$ 

This prevents  $\boldsymbol{U}\boldsymbol{V}^{\top}$  from being too peaky

• 
$$\mu_1 \leq \mu_2 \leq \mu_1^2 r$$
, since

$$\begin{split} |(\boldsymbol{U}\boldsymbol{V}^{\top})_{ij}| &= |\boldsymbol{e}_i^{\top}\boldsymbol{U}\boldsymbol{V}^{\top}\boldsymbol{e}_j| \leq \|\boldsymbol{e}_i^{\top}\boldsymbol{U}\|_2 \cdot \|\boldsymbol{V}^{\top}\boldsymbol{e}_j\|_2 \leq \frac{\mu_1 r}{n} \\ \|\boldsymbol{U}\boldsymbol{V}^{\top}\|_{\infty}^2 \geq \frac{\|\boldsymbol{U}\boldsymbol{V}^{\top}\boldsymbol{e}_j\|_{\mathrm{F}}^2}{n} = \frac{\|\boldsymbol{V}^{\top}\boldsymbol{e}_j\|_2^2}{n} = \frac{\mu_1 r}{n^2} \; (\text{suppose } \|\boldsymbol{V}^{\top}\boldsymbol{e}_j\|_2^2 = \frac{\mu_1 r}{n}) \end{split}$$

 $\begin{array}{ll} \mathsf{minimize}_{\boldsymbol{L},\boldsymbol{S}} & \mathsf{rank}(\boldsymbol{L}) + \lambda \|\boldsymbol{S}\|_0 & \mathsf{s.t.} & \boldsymbol{M} = \boldsymbol{L} + \boldsymbol{S} & (14.1) \\ & & & \downarrow \\ & & \\ \mathsf{minimize}_{\boldsymbol{L},\boldsymbol{S}} & \|\boldsymbol{L}\|_* + \lambda \|\boldsymbol{S}\|_1 & \mathsf{s.t.} & \boldsymbol{M} = \boldsymbol{L} + \boldsymbol{S} & (14.2) \end{array}$ 

- $\|\cdot\|_*$ : nuclear norm;  $\|\cdot\|_1$ : entry-wise  $\ell_1$  norm
- $\lambda > 0$ : regularization parameter that balances two terms

### Theorem 14.4 (Candès, Li, Ma, Wright '11)

- $\operatorname{rank}(L) \lesssim \frac{n}{\max\{\mu_1, \mu_2\} \log^2 n};$
- Nonzero entries of S are randomly located, and ||S||<sub>0</sub> ≤ ρ<sub>s</sub>n<sup>2</sup> for some constant ρ<sub>s</sub> > 0 (e.g. ρ<sub>s</sub> = 0.2).

Then (14.2) with  $\lambda = 1/\sqrt{n}$  is exact with high prob.

- rank(L) can be quite high (up to n/polylog(n))
- Parameter free:  $\lambda = 1/\sqrt{n}$
- Ability to correct gross error:  $\| \boldsymbol{S} \|_0 symp n^2$
- Sparse component S can have arbitrary magnitudes / signs!

## Geometry



#### Fig. credit: Candès '14

### **Empirical success rate**



Fig. credit: Candès, Li, Ma, Wright '11

Theorem 14.5 (Ganesh, Wright, Li, Candès, Ma'10, Chen, Jalali, Sanghavi, Caramanis'13)

• 
$$\operatorname{rank}(\boldsymbol{L}) \lesssim \frac{n}{\max\{\mu_1, \mu_2\} \log^2 n};$$

nor

 Nonzero entries of S are randomly located, have random sign, and ||S||<sub>0</sub> = ρ<sub>s</sub>n<sup>2</sup>.

Then (14.2) with  $\lambda \asymp \sqrt{rac{1ho_s}{
ho_s n}}$  succeeds with high prob., provided that

$$\underbrace{1-\rho_s}_{\text{p-corruption rate}} \gtrsim \sqrt{\frac{\max\{\mu_1,\mu_2\}r\operatorname{polylog}(n)}{n}}$$

• When additive corruptions have random signs, (14.2) works even when a dominant fraction of the entries are corrupted

- Matrix completion: does not need  $\mu_2$
- Robust PCA: so far we need  $\mu_2$

**Question:** is  $\mu_2$  needed? can we recover L with rank up to  $\frac{n}{\mu_1 \operatorname{polylog}(n)}$  (rather than  $\frac{n}{\max\{\mu_1, \mu_2\}\operatorname{polylog}(n)}$ )?

**Answer:** no (example: planted clique)

**Setup:** a graph  $\mathcal{G}$  of n nodes generated as follows

- 1. connect each pair of nodes independently with prob. 0.5
- 2. pick  $n_0$  nodes and make them a clique (fully connected)

**Goal:** find the hidden clique from  ${\mathcal{G}}$ 

Information theoretically, one can recover the clique if  $n_0 > 2 \log_2 n$ 

**Conjecture:**  $\forall$  constant  $\epsilon > 0$ , if  $n_0 \leq n^{0.5-\epsilon}$ , then no tractable algorithm can find the clique from  $\mathcal{G}$  with prob. 1 - o(1)

- often used as a hardness assumption

#### Lemma 14.6

If there is an algorithm that allows recovery of any L from M with  $\mathsf{rank}(L) \leq \frac{n}{\mu_1 \mathsf{polylog}(n)}$ , then the above conjecture is violated

Suppose L is the true adjacency matrix,

$$L_{i,j} = \begin{cases} 1, & \text{if } i, j \text{ are both in the clique} \\ 0, & \text{else} \end{cases}$$

Let A be the adjacency matrix of  $\mathcal{G}$ , and generate M s.t.

$$M_{i,j} = \begin{cases} A_{i,j}, & \text{with prob. } 2/3\\ 0, & \text{else} \end{cases}$$

Therefore, one can write

$$M = L +$$
  $M - L$ 

each entry is nonzero w.p. 1/3

Note that

$$\mu_1=rac{n}{n_0}$$
 and  $\mu_2=rac{n^2}{n_0^2}$ 

If there is an algorithm that can recover any  $\pmb{L}$  of rank  $\frac{n}{\mu_1 \mathrm{polylog}(n)}$  from  $\pmb{M},$  then

$$\mathsf{rank}(\boldsymbol{L}) = 1 \leq \frac{n}{\mu_1 \mathsf{polylog}(n)} \quad \Longleftrightarrow \quad n_0 \geq \mathsf{polylog}(n)$$

But this contradicts the conjecture (which claims computational infeasibility to recover L unless  $n_0 \ge n^{0.5-o(1)}$ )

What if we have missing data + corruptions?

• Observed entries

$$M_{ij} = L_{ij} + S_{ij}, \quad (i,j) \in \Omega$$

for some observation set  $\Omega$ , where  $\boldsymbol{S} = (S_{ij})$  is sparse

• A natural extension of RPCA

 $\text{minimize}_{\boldsymbol{L},\boldsymbol{S}} \quad \|\boldsymbol{L}\|_* + \lambda \|\boldsymbol{S}\|_1 \quad \text{s.t.} \ \mathcal{P}_{\Omega}(\boldsymbol{M}) = \mathcal{P}_{\Omega}(\boldsymbol{L} + \boldsymbol{S})$ 

• Theorems 14.4 - 14.5 easily extend to this setting

In the presence of noise, one needs to solve

$$\mathsf{minimize}_{oldsymbol{L},oldsymbol{S}} \hspace{0.1 in \|oldsymbol{L}\|_{*}} + \lambda \|oldsymbol{S}\|_{1} + rac{\mu}{2} \|oldsymbol{M} - oldsymbol{L} - oldsymbol{S}\|_{\mathrm{F}}^{2}$$

which can be solved efficiently

Algorithm 14.1 Iterative soft-thresholding

for  $t = 0, 1, \cdots$ :

$$egin{aligned} oldsymbol{L}^{t+1} &= \mathcal{T}_{1/\mu} \left(oldsymbol{M} - oldsymbol{S}^t 
ight) \ oldsymbol{S}^{t+1} &= \psi_{\lambda/\mu} \left(oldsymbol{M} - oldsymbol{L}^{t+1} 
ight) \end{aligned}$$

where  $\mathcal{T}:$  singular-value thresholding operator;  $\psi:$  soft thresholding operator

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